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# Uprooting and Rerooting Graphical Models

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## Abstract

We show how any binary pairwise model may be ‘uprooted’ to a fully symmetric model, wherein original singleton potentials are transformed to potentials on edges to an added variable, and then ‘rerooted’ to a new model on the original number of variables. The new model is essentially equivalent to the original model, with the same partition function and allowing recovery of the original marginals or a MAP configuration, yet may have very different computational properties that allow much more efficient inference. This meta-approach deepens our understanding, may be applied to any existing algorithm to yield improved methods in practice, generalizes earlier theoretical results, and reveals a remarkable interpretation of the triplet-consistent polytope.

## 1. Introduction

Undirected graphical models, also called Markov random fields (MRFs), have become a central tool in machine learning, providing a powerful and compact way to model relationships between variables. However, many key problems, such as identifying a configuration with highest probability (termed maximum a posteriori or MAP inference), estimating marginal probabilities of subsets of variables (marginal inference) or calculating the normalizing partition function, are typically computationally intractable, leading to much work to identify settings where exact polynomial-time methods apply, or to create approximate algorithms that perform well.

Focusing on binary pairwise models (see §2 for definitions and details), we provide a general meta-method for inference that generalizes and strengthens existing theoretical results, deepens our understanding, and can help significantly in practice. Suppose a model  $M$  has  $n$  variables  $X_1, \dots, X_n$  with various singleton and edge potentials. We start by *uprooting* this model to a uniquely determined ‘par-

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ent’ model  $M^+$  where all previous singleton potentials are converted to edge potentials to a newly added variable  $X_0$ . This  $M^+$  model is elegantly symmetric: no singleton potentials, and all edge potentials give a score only if incident variables are different. This uprooting is not a novel idea for MAP inference (Barahona et al., 1988; Sontag, 2007) but we believe the other ideas presented here are new.

The uprooted  $M^+$  model is interesting in itself; for example, its partition function is exactly twice that of the original model  $M$ , which we may consider as the parent  $M^+$  model *rooted* at  $X_0$ . A key idea is that we can *reroot*  $M^+$  at any other variable, for example  $X_2$ , to yield an equivalent model on  $n$  variables  $X_0, X_1, X_3, \dots, X_n$ , which has new singleton potentials determined by the edge potentials to  $X_2$  in  $M^+$ . In effect, this is a different view or ‘cristallization’ of the parent symmetric  $M^+$  model. Call this  $X_2$ -rooted model  $M_2$  (we could root at any variable  $X_i$  to obtain  $M_i$ ; note that the original model  $M$  is  $M_0$ ).

We make the following observations.

- $M_2$  indeed represents essentially the same model as  $M$ . It lies in the equivalence class of models that map to the same unique symmetric representation  $M^+$ .
- $M_2$  has the same partition function as  $M$  but may have very different computational properties. The original model  $M$  might be hard but  $M_2$  could be easy.
- Using any existing inference method for  $M_2$ , it is not hard also to recover all the original singleton marginals or a MAP configuration of  $M$ , see §4.1.
- Hence we have a general meta-method for inference: given any inference approach, instead of applying it to  $M$ , we can instead consider various equivalent rerooted models and apply the approach to one of them instead, which may work much better.
- Many existing methods and bounds apply only to particular ranges of edge and singleton potentials, which are changed after rerooting. Hence, we can generalize existing approaches. We discuss various implications in §5. For example, we can use the very efficient max flow/min cut method for MAP inference in a model if all edges are attractive with no conditions on singleton potentials (more generally if the model is *balanced*, see §2.1). This might not be possible in the original model  $M$  but will

be possible in some rerooted model iff there exists some variable  $X_i$  in  $M^+$  s.t. after rooting at  $X_i$ , the remainder  $M_i$  is balanced. This is equivalent to the condition that  $M^+$  is *almost balanced*. This can be tested efficiently.

- Understanding that singleton and pairwise potentials appear different only due to a particular choice of root in  $M^+$  provides an important fresh perspective, leading to a re-evaluation of existing methods, and a remarkable interpretation of the triplet-consistent polytope, see §5.

Binary pairwise models play an important role in many fields such as computer vision (Blake et al., 2011). Further, any discrete graphical model may essentially be converted into an equivalent binary pairwise model, though this may require a large increase in the number of variables.<sup>1</sup>

**Contributions.** After providing background in §2, we explain the details of the uprooting and rerooting approach in §3-4, showing how inference on a rerooted model allows recovery of information about the original model. This includes a discussion in §4.2 of the relation to clamping, where we introduce a new clamping heuristic that performs particularly well in settings that are likely to arise for rerooting. In §5, we discuss implications of rerooting, showing how it generalizes theoretical results, may be used as a meta-algorithm for inference methods, and provides a fascinating perspective on the triplet-consistent polytope. We provide an empirical evaluation in §6, showing that rerooting can be particularly effective for models with dense, strong edges and weak singleton potentials.

**Related Work.** What we call uprooting has been described previously as a way to reduce MAP inference of  $M$  to the MAXCUT problem in  $M^+$  (Barahona et al., 1988). As we discuss in §4.2, uprooting to  $M^+$  may be viewed as a de-clamping of the model at  $X_0$ , while a rerooting may be considered a re-clamping at a different variable. Hence, rerooting replaces an initial implicit clamp choice at  $X_0$  with another. The choice of which root to choose is essentially the question of which variable in  $M^+$  to clamp. Methods to select a variable to clamp have been explored by Eaton and Ghahramani (2009) and Weller and Domke (2016).

## 2. Preliminaries

We focus on undirected graphical models with  $n$  binary variables  $X_1, \dots, X_n \in \{0, 1\}$ . We consider a probability distribution  $p(x) = e^{\theta(x)}/Z(\theta)$  where  $x = (x_1, \dots, x_n)$  is one particular configuration of all variables and  $\theta(x)$  is the score of configuration  $x$ , which decomposes into singleton and pairwise (log) potential terms. The topology of

<sup>1</sup>Eaton and Ghahramani (2013) show that this is strictly true if all model states have probability strictly  $> 0$ , otherwise an arbitrarily good approximation is possible.

the model is the graph  $(\mathcal{V}, \mathcal{E})$ , that is  $\mathcal{V} = \{1, \dots, n\}$  where  $i$  corresponds to  $X_i$ , and  $\mathcal{E} \subseteq \mathcal{V} \times \mathcal{V}$  contains an edge for each pairwise score relationship. We assume a reparameterization to the minimal representation (Wainwright and Jordan, 2008) wherein the score may be written

$$\theta(x) = \sum_{i \in \mathcal{V}} \theta_i x_i - \frac{1}{2} \sum_{(i,j) \in \mathcal{E}} W_{ij} \mathbb{1}[x_i \neq x_j], \quad (1)$$

where  $\mathbb{1}[\cdot]$  is the indicator function.  $Z(\theta)$  is a normalizing constant, called the *partition function*, which ensures that probabilities sum to 1, i.e.  $Z(\theta) = \sum_{x \in \{0,1\}^n} e^{\theta(x)}$ .

Note that (1) gives a score to an edge only if its incident variables are different. The factor of  $-\frac{1}{2}$  before the edge potentials means that the signs and scaling of our parameters are consistent with earlier work such as (Welling and Teh, 2001; Weller and Domke, 2016). If  $W_{ij} > 0$  then the edge  $(i, j)$  is *attractive* and tends to pull its incident variables towards the same value; if  $W_{ij} < 0$  then the edge is *repulsive* and tends to push apart the values of its variables.

### 2.1. Attractive, Mixed and Balanced Models

If all edges of a model are attractive, i.e. if  $W_{ij} \geq 0 \forall (i, j) \in \mathcal{E}$ , then we say that the model is attractive, else it is mixed. Sometimes a mixed model may be converted to an equivalent attractive model by flipping (also called switching) a subset of variables  $S$ , which reverses the signs of their singleton potentials and of the edge potentials between variables in  $S$  and variables in  $\mathcal{V} \setminus S$ ; if possible, such a mixed model is called *balanced*. Harary (1953) showed that a model is balanced iff it does not contain a *frustrated cycle*, which is a cycle containing an odd number of repulsive edges. This may be checked and, if balanced, then a flipping set  $S$  found, in time linear in  $|\mathcal{E}|$  (Harary and Kabell, 1980). Hence, results for attractive models readily extend to the larger class of balanced models.

**Notation.** For any  $a \in \{0, 1\}$ , let  $\bar{a} = 1 - a$  (this flips  $0 \leftrightarrow 1$ ). Similarly, for a vector  $x = (x_1, \dots, x_n)$ , let  $\bar{x} = (\bar{x}_1, \dots, \bar{x}_n) = (1 - x_1, \dots, 1 - x_n)$ . For a configuration  $y = (y_0, y_1, \dots, y_n)$  of  $M^+$ , and  $a \in \{0, 1\}$ , we may write  $y = (a, x)$  to mean  $y = (a, x_1, \dots, x_n)$ .

## 3. Uprooting a Model

We show how any model  $M$  on  $n$  variables  $X_1, \dots, X_n$  with singleton potentials may be uprooted to a unique symmetric (i.e. no singleton potentials) model  $M^+$  on  $n + 1$  variables  $X_0, X_1, \dots, X_n$ . Edges to the extra variable  $X_0$  encode the original singleton potentials.<sup>2</sup>

<sup>2</sup>If the original model  $M$  has no singleton potentials, then it may be regarded as already in  $M^+$  form. It may still be rooted at any variable, as described in §4.

M <sup>+</sup> configuration				edges: score ✓ if ends are different					
x <sub>0</sub>	x <sub>1</sub>	x <sub>2</sub>	x <sub>3</sub>	e <sub>01</sub>	e <sub>02</sub>	e <sub>03</sub>	e <sub>12</sub>	e <sub>13</sub>	e <sub>23</sub>
0	0	0	0						
<b>0</b>	<b>0</b>	<b>0</b>	<b>1</b>		✓		✓	✓	
0	0	1	0		✓		✓		✓
0	0	1	<b>1</b>		✓	✓	✓	✓	
0	1	0	0	✓			✓	✓	
<b>0</b>	<b>1</b>	<b>0</b>	<b>1</b>	✓		✓	✓		✓
0	1	1	0	✓	✓			✓	✓
0	1	1	<b>1</b>	✓	✓	✓			
<b>1</b>	<b>0</b>	<b>0</b>	<b>0</b>	✓	✓	✓			
1	0	0	1	✓	✓			✓	✓
1	0	1	0	✓		✓	✓		✓
1	0	1	<b>1</b>	✓		✓	✓		
<b>1</b>	<b>1</b>	<b>0</b>	<b>0</b>		✓	✓	✓	✓	
1	1	0	1		✓		✓		✓
1	1	1	0			✓		✓	✓
1	1	1	1						

Table 1. An example showing all configurations of an uprooted M<sup>+</sup> model on 4 variables. The original model M has 3 variables X<sub>1</sub>, X<sub>2</sub>, X<sub>3</sub> then X<sub>0</sub> was added. Each configuration of M features twice: once as (0, x<sub>1</sub>, x<sub>2</sub>, x<sub>3</sub>) in the top half, and then again with all settings flipped as (1, x̄<sub>1</sub>, x̄<sub>2</sub>, x̄<sub>3</sub>) in the bottom. Each has the same score in M<sup>+</sup>, with the score determined only by the edges which are activated: see the check marks on the right and note their reflective symmetry across the horizontal line in the middle of the table.

The pink shaded rows indicate the configurations for the rerooted model M<sub>2</sub> where X<sub>2</sub> = 0. Observe that given the symmetry, these correspond 1-1 with the configurations of M. Hence, we can recover the partition function, marginal probabilities or a MAP configuration for M by inference on M<sub>2</sub>. For example, p<sub>0</sub>(X<sub>3</sub> = 1) for M may be computed as p<sub>2</sub>(X<sub>3</sub> ≠ X<sub>0</sub>) for M<sub>2</sub>, i.e. sum over the rows shown in bold. Each of the rows in the top half with x<sub>3</sub> = 1 which is missing from M<sub>2</sub> (that is, not shaded pink) has an exactly corresponding row in the bottom half, as indicated by blue curves in the table. See §3 and §4 for details.

Let  $y = (y_0, y_1, \dots, y_n)$  be a configuration in M<sup>+</sup> of its  $n+1$  variables, and let  $\phi(y)$  be its score in M<sup>+</sup>. Requiring  $\phi(y)$  to be in the same form as (1) but with no singleton potentials, and to match the scores of configurations in M when  $x_0 = 0$ , i.e. requiring  $\phi(0, x) = \theta(x)$ , implies

$$\phi(y) = -\frac{1}{2} \sum_{\mathcal{E}'} W_{ij} \mathbf{1}[y_i \neq y_j], \quad (2)$$

where the edges of M<sup>+</sup> are  $\mathcal{E}' = \mathcal{E} \cup \mathcal{F}$  consisting of the original edges  $\mathcal{E}$  of M, together with new edges  $\mathcal{F}$  which are added to the new variable X<sub>0</sub>, given by  $\mathcal{F} = \{(0, i) : \theta_i \neq 0\}$ . Weights for edges in  $\mathcal{E}$  remain unchanged. Weights for the new edges in  $\mathcal{F}$  are set as  $W_{0i} = -2\theta_i$ . To see this, note that the singleton potentials in (1) are already in the form  $\theta_i \mathbf{1}[x_i \neq x_0 | x_0 = 0]$ .

Note the sign flip when a singleton potential is converted to

an edge potential, e.g.  $\theta_i > 0$  becomes a repulsive edge in M<sup>+</sup> with  $W_{0i} < 0$ . This is an unavoidable consequence of choosing parameters in (1) to match earlier work. It may be helpful to think of  $\theta_i > 0$  as encouraging  $X_i$  to be different to 0, i.e. a repulsive edge from  $X_0 = 0$ .

Observe that each configuration  $x$  of M maps to *two* configurations  $y_0$  and  $y_1$  in M<sup>+</sup>,

$$M: x = (x_1, \dots, x_n) \rightarrow M^+: \begin{cases} y_0 = (0, x) \\ y_1 = \bar{y}_0 = (1, \bar{x}), \end{cases} \quad (3)$$

i.e.  $y_0 = (0, x_1, \dots, x_n)$  and  $y_1 = \bar{y}_0 = (1, \bar{x}_1, \dots, \bar{x}_n)$ . Given the symmetry of (2), it is clear that  $\phi(y_0) = \phi(y_1)$ . See Table 1 and Figure 1 for an example.

The partition function for M<sup>+</sup> is clearly twice that of M, i.e.  $Z(\phi) = 2Z(\theta)$ . The original model M is exactly M<sup>+</sup> conditioned on  $X_0 = 0$ , and we can write  $M = M_0$ .

## 4. Rerooting a Model

The symmetric model M<sup>+</sup> with form (2) may be rooted at any variable  $X_i$  by conditioning on  $X_i = 0$  to yield a model on  $n$  variables consisting of  $\{X_0, \dots, X_{i-1}, X_{i+1}, \dots, X_n\}$ , which we write as  $M_i$ .<sup>3</sup> See Table 1 and Figure 1 for an example.

Considering (3), for any  $i$ , there is a score-preserving 1-1 correspondence between configurations in M and those in  $M_i$  which matches  $x$  in M with whichever of  $y_0$  or  $y_1$  has  $x_i = 0$  (the  $x_i$  coordinate is removed to give the configuration in  $M_i$ ). Equivalently, if  $x$  in M has  $x_i = 0$ , then it matches to  $(0, x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n)$  in  $M_i$ , otherwise it matches to the same configuration but with all settings flipped, i.e.  $(1, \bar{x}_1, \dots, \bar{x}_{i-1}, \bar{x}_{i+1}, \dots, \bar{x}_n)$ .

### 4.1. Recovery of Original MAP Configuration, Partition Function or Marginals

In this Section, we show that if inference can be performed on a rerooted model  $M_i$ , then we can recover results for the original model M.

#### 4.1.1. MAP INFERENCE

For MAP inference, given the score-preserving 1-1 correspondence of configurations noted above in §4, if a MAP configuration is determined for  $M_i$ , then the corresponding configuration in M is a MAP configuration for M with the same score. Specifically, we have the following result.

**Lemma 1.** *If  $(x_0^*, \dots, x_{i-1}^*, x_{i+1}^*, \dots, x_n^*)$  is a MAP con-*

<sup>3</sup>Given the symmetry, one could instead equivalently consider M<sup>+</sup> conditioned on  $X_i = 1$  but then one would need to flip variables after performing inference in order to match the original interpretation in M.

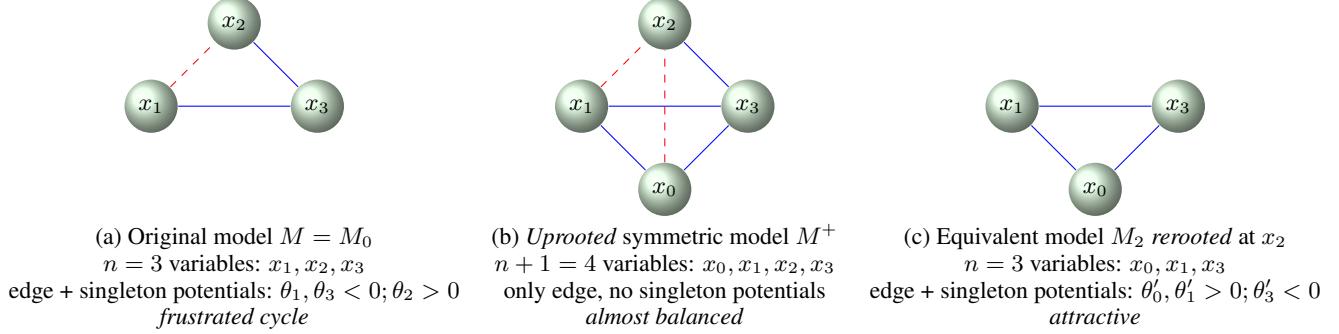


Figure 1. Examples of (a) an original model  $M = M_0$  on three variables, together with (b) its unique uprooted model  $M^+$ , and (c) a different rooting of  $M^+$  at  $x_2$  to yield  $M_2$ , where now all edges are attractive. Solid blue (dashed red) edges are attractive (repulsive).

figuration for  $M_i$ , then the corresponding MAP configuration for  $M$ , with the same score, is:

$$\begin{cases} m = (x_1^*, \dots, x_{i-1}^*, x_i = 0, x_{i+1}^*, \dots, x_n^*) & \text{if } x_0^* = 0 \\ \bar{m} = (\bar{x}_1^*, \dots, \bar{x}_{i-1}^*, x_i = 1, \bar{x}_{i+1}^*, \dots, \bar{x}_n^*) & \text{if } x_0^* = 1. \end{cases}$$

#### 4.1.2. MARGINAL INFERENCE AND COMPUTING Z

Since  $M_i$  and  $M$  have corresponding configurations with equal scores, they have the same partition function.

In order to differentiate between probabilities obtained for different models, we use the following notation: let  $p_i$  be the probability distribution in model  $M_i$ , in particular  $p_0$  is the distribution for model  $M_0$  which is the original model  $M$ ; let  $p_+$  be the distribution in the uprooted model  $M^+$ .

Each model  $M_i$  is the result of conditioning on  $X_i = 0$  in  $M^+$ . We would like to perform (exact or approximate) inference on  $M_i$  to obtain  $p_i$ , then use this to recover marginals  $p_0(X_j = 1) \forall j \in \{1, \dots, n\}$ . The following result achieves this. See Table 1 for an example.

**Lemma 2.** Given distribution  $p_i$  for any  $i \in \{1, \dots, n\}$ , the marginals  $p_0(X_j = 1)$  for the original model  $M = M_0$  may be recovered as follows:

$$p_0(X_j = 1) = \begin{cases} p_i(X_0 = 1) & j = i \\ p_i(X_j \neq X_0) & j \neq i. \end{cases}$$

*Proof.* This follows from the symmetry of  $M^+$ , see the Appendix for details.  $\square$

#### 4.2. Relation to Clamping, How to Choose a Root?

Conditioning a model on a variable taking a particular value is sometimes called *clamping* (Eaton and Ghahramani, 2009; Weller and Jebara, 2014). Since  $M_i$  is  $M^+$  conditioned on  $X_i = 0$ , we may view uprooting from  $M = M_0$  to  $M^+$  as *de-clamping*  $X_0$  back to a parent model; then rerooting at variable  $X_i$  is a *re-clamping* at  $X_i = 0$  to obtain  $M_i$ .

In typical clamping for inference, one must condition a variable to each of its values and combine all results (for example, if estimating  $Z$ , one must sum the approximate subpartition functions). For binary variables, this requires running your inference algorithm twice. In contrast, a rooted model gets a ‘clamping for free’ at the root variable, with just one inference run required.

A natural question is how to choose a good root when rerooting a model? Given the interpretation of rooting as clamping, we can draw on earlier work. Weller and Domke (2016) examined a range of heuristics and concluded that a fast method called maxW typically performs very well for approximate inference.<sup>4</sup> The idea behind maxW is that many popular methods of approximate inference, such as belief propagation, are exact on acyclic models but can perform poorly when there are cycles composed of strong edge weights. It is NP-hard to identify heavy cycles but the following simple heuristic was shown to be empirically effective. For each variable, a sum of absolute values of incident edge weights is computed, then the variable with the highest sum is selected to clamp. When it is clamped, this variable is effectively removed from the model, thereby eliminating any cycles which ran through it.

#### 4.2.1. A NEW METHOD: MAXTW

In §6, we explore the value of rerooting using the earlier maxW heuristic to select the root variable. We observe that maxW sometimes performs well, but one setting where it can perform poorly is if a choice must be made between one variable that has a few strong edges and another that has many weak edges. When rerooting, this may happen frequently. For example, consider an initial model  $M$  with a grid topology, and singleton potentials that are low relative to edge potentials: in  $M^+$  this leads to  $X_0$  having a weak

<sup>4</sup>Weller and Domke (2016) showed that a more sophisticated variant called maxW+core+TRE performed slightly better in general, but TRE is redundant for the fully symmetric  $M^+$  model, and the core idea makes no difference in the experiments we run.

edge to every other variable, whereas other variables have few strong edges.  $\text{maxW}$  simply picks the variable  $i$  with highest  $\sum_{j \in \mathcal{N}(i)} |W_{ij}|$ , where  $\mathcal{N}(i)$  is the set of variables adjacent to  $i$ . However, the direct influence of a strong edge weight does not keep increasing linearly with its weight, rather it reaches a hard saturation level (Weller and Jebara, 2014, Supplement, Lemma 12). Here we address this by introducing an alternative heuristic we call  $\text{maxtW}$ , which picks the variable  $i$  with  $\max \sum_{j \in \mathcal{N}(i)} \tanh \left| \frac{W_{ij}}{4} \right|$ . This form was chosen based on earlier work on loop series expansions (Weller et al., 2014; Sudderth et al., 2007). See results in §6 where  $\text{maxtW}$  can dramatically outperform  $\text{maxW}$  on grids.

## 5. Implications of Rerooting

Rerooting provides a conceptual framework to view singleton and edge potentials as essentially equivalent except for a choice of rooting of the symmetric uprooted  $M^+$  parent model. After rerooting, it may be possible to apply many methods or bounds that were unavailable for the original model  $M$ . We consider important examples below.

### 5.1. MAP Inference

The success of many existing methods of MAP inference depends critically on the nature of the edge potentials of a model, but can be relatively insensitive to the singleton potentials. For example, both the max flow/min cut method (Greig et al., 1989) and the basic linear programming (LP) relaxation over the local polytope LOC (Wainwright and Jordan, 2008) provide an exact solution in polynomial-time if the model is attractive. These approaches generalize to balanced models, see §2.1.

With rerooting, these methods can now be used on the significantly larger class of model where some rooting  $M_i$  exists which is balanced. This holds iff the uprooted model  $M^+$  is *almost balanced*, which means it contains a variable such that removing it renders the remaining model balanced. See Figure 1 for an example.

Almost balanced models have received recent attention. Jebara (2009) introduced a method for MAP inference via a reduction to the graph-theoretic challenge of identifying a *maximum weight stable set* (MWSS) in a derived weighted graph, which if *perfect*, allows an exact solution to be obtained efficiently. Weller (2015b) showed that this method applies iff the *block decomposition* of the model  $M$  yields blocks (maximal 2-connected components) which are all almost balanced. With rerooting, we can extend this method to models  $M$  that have uprooted models  $M^+$  that are *2-almost balanced*, i.e. models which can be rendered balanced by deleting 2 variables (since by rooting at either of these variables, the rooted model is almost balanced).

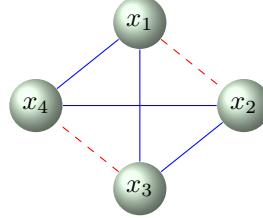


Figure 2. An example of a model  $M$  which is not almost balanced, hence does not satisfy the conditions of Weller et al. (2016) for tightness of LP+TRI. Nevertheless, by Theorem 4, it is sufficient if its uprooted  $M^+$  is 2-almost balanced. Hence LP+TRI will always be tight for (any rerooting of) this model provided singleton potentials are not of the form:  $\theta_1, \theta_2$  take one sign (either positive or negative); and  $\theta_3, \theta_4$  take the other. Solid blue (dashed red) edges are attractive (repulsive). See §5.2 for details.

### 5.2. Local and Triplet Polytopes, Why ‘Rooting’?

Weller et al. (2016) showed that the LP relaxation on the triplet-consistent polytope, LP+TRI, yields an exact MAP configuration of a model  $M$  provided it is almost balanced (for any singleton potentials). As above this can now be generalized to be used for any model if its uprooted model  $M^+$  is 2-almost balanced, since then a rooting exists which is almost balanced. In fact, we can achieve a much stronger result due to the following remarkable property of TRI.

**Theorem 3** (TRI is ‘universally rooted’). *LP+TRI yields the same optimum score for  $M$  as for any rerooting  $M_i$ ; hence LP+TRI is either tight for all rerootings or for none.*

Theorem 3 immediately yields the following new result.

**Theorem 4.** *LP+TRI is tight for (any rerooting of) a model  $M$  whose uprooted model  $M^+$  is 2-almost balanced.*

See the Appendix §8 for details and proofs. This beautifully shows the common nature of edge and singleton potentials for TRI, examining the signs of all edges in  $M^+$  in the same symmetric way.

Theorem 4 helps us to understand tightness of LP+TRI on real-world vision tasks, where learned models are close to attractive due to contiguity of objects. As a small example, Theorem 4 shows that LP+TRI is tight for the model shown in Figure 2, despite it not being almost balanced, provided the signs of singleton potentials leave  $M^+$  2-almost balanced (this holds for all values unless:  $\theta_1, \theta_2$  take one sign (positive or negative); and  $\theta_3, \theta_4$  take the other).

Here we sketch the reasoning. The following polytopes are equivalent (see Deza and Laurent, 1997):

$$\begin{array}{ccc}
 n \text{ variables} + \binom{n}{2} \text{ edges} & & \binom{n+1}{2} \text{ edges} \\
 \text{Marginal polytope of } M & \leftrightarrow & \text{Cut polytope of } M^+ \\
 \text{TRI relaxation} & \leftrightarrow & \text{MET relaxation} \\
 \text{LOC relaxation} & \leftrightarrow & \text{RMET relaxation}
 \end{array}$$

MET, the semimetric polytope relaxation of the cut

polytope, enforces triplet constraints on *every triplet* of  $\{X_0, \dots, X_n\}$ . In contrast, RMET, the *rooted* semimetric polytope, enforces these same triplet constraints *only on triplets that include the root  $X_0$  variable*.

This explains the name *rooting*. LOC is equivalent to a specifically rooted RMET polytope, which is why approaches over LOC (including many message-passing algorithms) deal differently with singleton and edge potentials, and might benefit significantly from rerooting. TRI, however, is equivalent to MET, which deals symmetrically with all variables in  $M^+$  and corresponds to *a simultaneous rooting at every variable*. This intriguing observation likely has further theoretical and algorithmic consequences, which we aim to explore in future work.

### 5.3. Belief Propagation

Belief propagation (BP, Pearl, 1988), or more generally the Bethe approximation (Yedidia et al., 2000), is a widely used approach for approximate inference, guaranteed to yield exact results in linear time for models without cycles. When applied to models with cycles, it often yields strikingly accurate results but may fail to converge altogether.

Much work has analyzed the convergence of BP, and the uniqueness of a fixed point, relying either on the strength of edge interactions (Heskes, 2004; Mooij and Kappen, 2005), or just on their signs (Watanabe, 2011). Mooij and Kappen (2007) refined their earlier result by considering also singleton potentials, but these are incorporated quite differently to the edge potentials. Hence, by rerooting it may be possible to provide theoretical guarantees on performance that are not available on the initial model.

As one example, it is known that if a model has one cycle, then the Bethe free energy is convex and BP has a unique fixed point (Pakzad and Anantharam, 2002). Consider the model shown in Figure 1. The original model (a) is a frustrated cycle, hence the BP estimate of  $Z$  will be too high, with unbounded high error as edge weights increase (Weller, 2015a, §6.3). In contrast, the rerooted model (c) is attractive, hence the BP estimate is always in the range  $[Z/2, Z]$  for any potentials (Weller and Jebara, 2014).

### 5.4. FPRAS, Bounds

Jerrum and Sinclair (1993) devised a *fully polynomial-time randomized approximation scheme* (FPRAS) for the partition function of a model  $M$  provided it is attractive and all singleton potentials are consistent in taking the same sign (positive or negative). This generalizes to any model with uprooted model  $M^+$  which is balanced (see §2.1).

Various methods have been developed to bound the partition function or marginals of a model (Leisink and Kappen, 2003; Ihler, 2007; Mooij and Kappen, 2008). These

treat singleton and edge potentials differently, hence may be generalized by considering rerootings.

### 5.5. Remarks

**Comparison to clamping.** Some of the benefits of rerooting could also be obtained by usual clamping of  $M$ . For example, if a model can be rendered balanced by rerooting at  $X_i$ , then this could also be achieved by clamping  $X_i$  in the original model. However, this would require performing multiple inference runs and combining results, rather than using the ‘free clamping’ available with a rerooting, see §4.2. Further, several results, including Theorem 4 and the observations in §5.2 on the triplet polytope, are not possible without considering rerooting.

**Evaluation of inference methods.** Approximate inference methods are typically evaluated empirically on a range of models, where singleton and edge potentials are treated quite differently. Often singleton potentials are drawn from some fixed narrow range while edge potentials are drawn from a range whose width is varied widely. From an uprooted model perspective, singleton and edge potentials are equivalent. Hence: (i) Varying singleton and edge potentials differently in empirical evaluations may be a peculiar assumption, though it could be justified as reflecting typical patterns in the real world; (ii) The implicit choice of root may be poor in some cases (i.e. results might be improved significantly by rerooting), which will obscure the underlying performance attributes of the inference method. We examine the extent of this effect in §6.

## 6. Experiments

Following the observation in §5.5, we are interested in the effect of rerooting in standard settings for empirical evaluation. We compared performance of estimating the partition function and singleton marginals after different rerootings of three popular approximation methods: Bethe (using the approach of Heskes et al., 2003 to ensure convergence), tree-reweighted (TRW, Wainwright et al., 2005) and naive mean field. For true values, we used the junction tree algorithm. All methods were implemented using libDAI (Mooij, 2010), see the Appendix §9 for details.

We ran experiments on the following topologies and model sizes: complete graphs on 10 and 15 variables; grids of size  $5 \times 5$  and  $9 \times 9$ . All potentials were drawn randomly: mixed models used  $W_{ij} \sim U[-W_{\max}, W_{\max}]$ , attractive models used  $W_{ij} \sim U[0, W_{\max}]$ , as  $W_{\max}$  was varied; singleton potentials were drawn either from a low range  $\theta_i \sim [-0.1, 0.1]$ , medium range  $\theta_i \sim [-2, 2]$ , or from a range commensurate with edge potentials, i.e.  $\theta_i \sim U[-W_{\max}/2, W_{\max}/2]$ , with the factor of 2 needed given the form of (1). These settings allow direct compar-

ison to earlier work such as by [Weller and Domke \(2016\)](#) or [Weller et al. \(2014\)](#). Others ([Meshi et al., 2009](#); [Sonsteg and Jaakkola, 2007](#)) use binary variables with values in  $\{-1, 1\}$  instead of  $\{0, 1\}$ , hence their edge (singleton) potentials should be multiplied by 4 (2, respectively) when making comparisons. We plot average error over 100 random runs for each setting. All results are in the Appendix.

As in §4.2, any rooting of a model  $M$  may be considered a clamping of the uprooted model  $M^+$ . The original model  $M = M_0$  implicitly reflects the decision to clamp at  $X_0$ , which might be a good or bad choice depending on the setting. Recall from §4.2 that  $\text{maxW}$  often performs well for selecting a variable to clamp, picking one with highest sum of incident edge strengths (taking absolute values). However, if a choice must be made between variable  $A$  with many weak edges, or  $B$  with few strong edges,  $\text{maxW}$  may make a poor choice by not recognizing that  $A$  is often better since the influence of strong edges saturates. Hence we introduced the  $\text{maxtW}$  heuristic in §4.2.1, which selects variable  $X_i$  with  $\max \sum_{j \in \mathcal{N}(i)} \tanh \left| \frac{W_{ij}}{4} \right|$ .

Our plots show average error when applying the approximate inference method to: the original model  $M$ ; the uprooted model  $M^+$ ; then rerootings at: the *worst* variable, the *best* variable, the *maxW* variable, and the *maxtW* variable. *Best* and *worst* always refer to the variable at which rerooting gave with hindsight the best and worst error for the partition function (even in plots for marginals).

## 6.1. Results

Figure 3 summarizes results for Bethe, typically the most accurate method. Looking across all results (see Appendix §9), we make the following observations.

For complete graphs,  $\text{maxW}$  and  $\text{maxtW}$  perform well. Rerooting is very effective as edge strength grows, both at low and medium levels of singleton potentials. This makes sense, since in this setting, the default rooting at  $X_0$  has relatively weak edges, and all variables in  $M^+$  have the same number of edges, hence it is likely to be very beneficial to switch to a different root with stronger edges. When singleton and edge potentials vary together, edges in  $M^+$  are all similar, but  $X_0$  is an average variable to clamp, whereas we do somewhat better by choosing a good variable.

For grids,  $\text{maxtW}$  is much better than  $\text{maxW}$  ( $\text{maxW}$  performs very poorly in some cases), appearing to handle uneven edge weights in  $M^+$  well. At low singleton potentials, rerooting is very helpful but this benefit disappears for stronger singleton potentials, where the original rooting performs equally to  $\text{maxtW}$ .

Results for MF and TRW are qualitatively similar to Bethe, with Bethe typically performing best. For mixed models with strong edges, MF performs very well. This is

likely due to MF optimizing within the marginal polytope, whereas Bethe and TRW use the local polytope, in which strong frustrated cycles can lead to high error.

Based on  $\text{maxtW}$ , we can suggest a guideline for when rerooting is likely to be helpful. For example, for a 4-way grid with  $n$  variables, constant singleton potentials  $T$  and edge weights  $W$ :  $4 \tanh \frac{W}{4} + \tanh \frac{2T}{4} > n \tanh \frac{2T}{4} \Leftrightarrow 4 \tanh \frac{W}{4} > (n-1) \tanh \frac{T}{2}$ . This is conservative since more randomness increases the value of rerooting by raising the chance of a better root. Demonstrating this, observe in Figure 3 that when singleton potentials are low, the improvement in  $\log Z$  estimate from rerooting using  $\text{maxtW}$  is about the same for  $9 \times 9$  grids as for smaller  $5 \times 5$  grids.

## 7. Conclusion

We introduced the idea of uprooting and then rerooting any binary pairwise graphical model. This immediately leads to a meta-algorithm for inference into which any existing approach may be slotted, and generalizes important theoretical results. Further, it provides an elegant conceptual framework for rethinking singleton and edge potentials with methodological consequences for how we evaluate models and methods. One application in §5.2 leads to Theorem 4, a strong result for tightness of LP relaxations on the triplet-consistent polytope TRI, and a remarkable interpretation of TRI as universally rooted.

Rerooting switches an implicit clamp choice in the uprooted model at  $X_0$  (perhaps a poor choice), instead to a carefully selected clamp choice, almost for free. This applies even for large models where it is desirable to clamp a series of variables: by rerooting, we may obtain one of the series for free, sometimes achieving dramatic improvements in accuracy with little work required. If there are multiple connected components, each should be handled separately, with its own  $X_0$ -type variable. This could be useful for (repeatedly) composing clamping and then rerooting each separated component.

Rerooting is particularly effective when a model has dense, strong edge weights and weak singleton potentials (a difficult setting for many existing methods). Our new  $\text{maxtW}$  heuristic performs particularly well in this setting (and should also be helpful for standard clamping approaches), sometimes dramatically outperforming the earlier  $\text{maxW}$  method.  $\text{maxtW}$  also provides a useful guideline for when uprooting is likely to be helpful, see the last paragraph of §6.1.

It will be interesting in future work to study further consequences of our interpretation of the triplet-consistent polytope, to consider the value of rerooting for approaches to learning graphical models, and to explore the benefits of rerooting when variables have a higher number of labels.

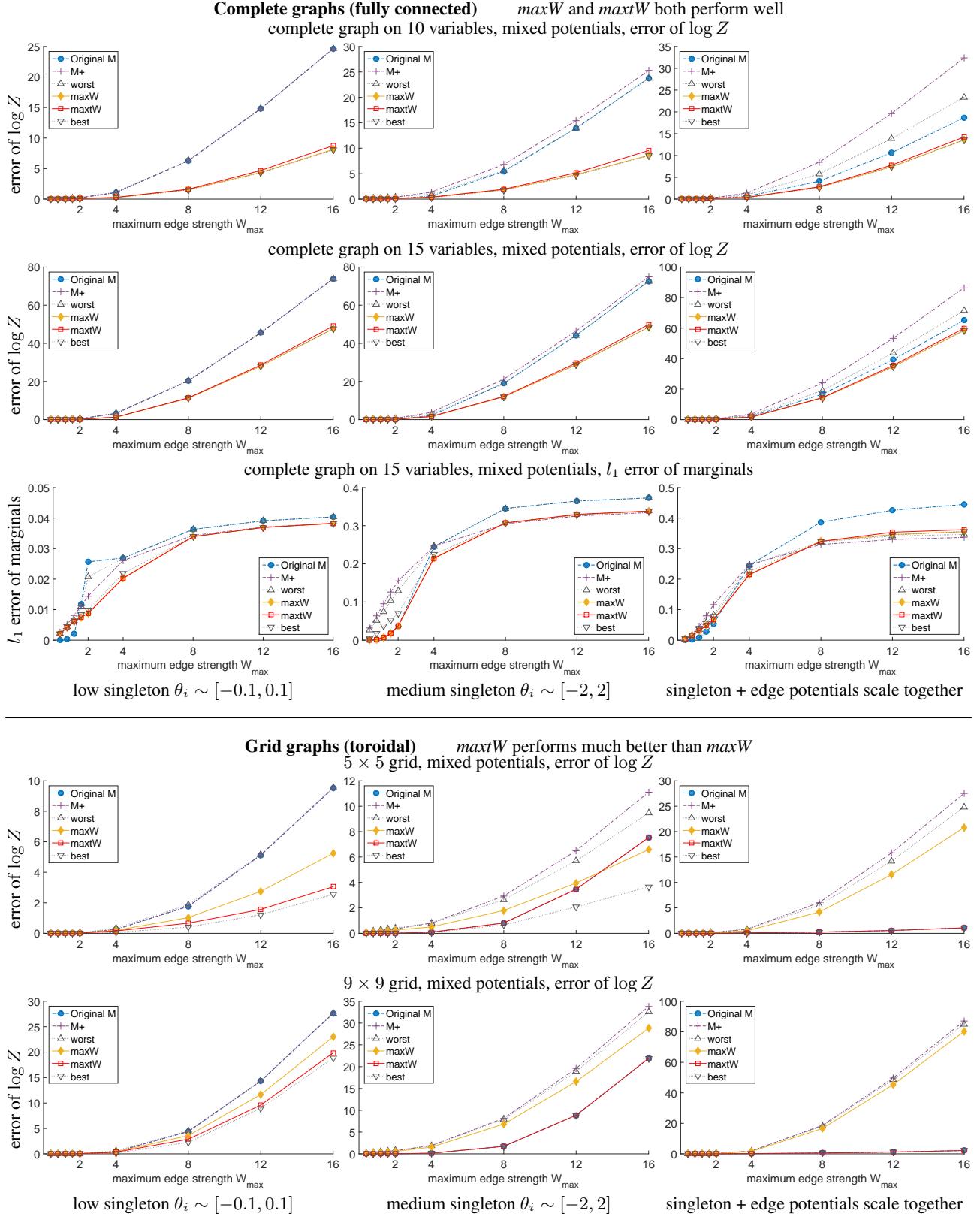


Figure 3. Average error of Bethe approximation for models with mixed potentials over 100 runs, showing smaller and larger models for comparison. Top: complete graphs (10 and 15 variables). Bottom: toroidal grid graphs (5  $\times$  5 and 9  $\times$  9). Each column shows different settings for singleton potentials: left is low range; centre is medium range; right varies singleton and edge potentials together. See §6.

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## APPENDIX: SUPPLEMENTARY MATERIAL

### Uprooting and Rerooting Graphical Models

Here we provide:

- Proof of Lemma 2.
- In §8, details of the polytopes and proofs from §5.2, including proofs of Theorems 3 and 4.
- In §9, details of experimental methods, and additional results.

**Lemma 2.** Given distribution  $p_i$  for any  $i \in \{1, \dots, n\}$ , the marginals  $p_0(X_j = 1)$  for the original model  $M = M_0$  may be recovered as follows:

$$p_0(X_j = 1) = \begin{cases} p_i(X_0 = 1) & j = i \\ p_i(X_j \neq X_0) & j \neq i. \end{cases}$$

*Proof.* First, for  $j = i$  we have

$$\begin{aligned} p_0(X_i = 1) &= p_+(X_i = 1 | X_0 = 0) \\ &= \frac{p_+(X_i = 1, X_0 = 0)}{p_+(X_0 = 0)} \\ &= \frac{p_+(X_i = 0, X_0 = 1)}{p_+(X_i = 0)} \quad (\text{symmetry of } M^+, \text{ note that } p_+(X_r = 0) = \frac{1}{2} \text{ for any } r \in \{0, \dots, n\}) \\ &= p_i(X_0 = 1). \end{aligned}$$

Next, for  $j \neq i$ , again using symmetry of  $M^+$ ,

$$\begin{aligned} p_0(X_j = 1) &= p_+(X_j = 1 | X_0 = 0) \\ &= \frac{p_+(X_j = 1, X_0 = 0)}{p_+(X_0 = 0)} \\ &= \frac{p_+(X_j = 1, X_0 = 0, X_i = 0) + p_+(X_j = 1, X_0 = 0, X_i = 1)}{p_+(X_0 = 0)} \\ &= \frac{p_+(X_j = 1, X_0 = 0, X_i = 0) + p_+(X_j = 0, X_0 = 1, X_i = 0)}{p_+(X_i = 0)} \\ &= \frac{p_+(X_j \neq X_0, X_i = 0)}{p_+(X_i = 0)} \\ &= p_i(X_j \neq X_0). \end{aligned}$$

□

## 8. Details of the polytopes and proofs from section 5.2

Weller et al. (2016) showed that LP+TRI is tight (that is, the LP relaxation on the triplet-consistent polytope is guaranteed to yield an optimum at an integral vertex) for any model which is almost balanced (that is, any model which contains a variable s.t. if it is removed then the remaining model is balanced; any singleton potentials are allowed). We first provide background and preliminary results in §8.1-8.2. For more extensive background, see (Wainwright and Jordan, 2008, Chapter 8), (Sontag, 2007) or (Deza and Laurent, 1997).

In §8.3, we prove Theorem 3, a general result which shows that TRI is ‘universally rooted’. Many optimization results that apply for TRI for *some* rerooting of a model will automatically apply for *all* rerootings.

We shall apply Theorem 3 to show how the result of Weller et al. (2016) may be significantly strengthened in Theorem 4 to demonstrate tightness of LP+TRI for any model  $M$  whose uprooted model  $M^+$  is 2-almost balanced (that is, the uprooted model contains 2 variables s.t. if they are both removed then what remains in the uprooted model is balanced).

**Notation.** As in 4.1.2, in order to differentiate between probabilities obtained for an initial model  $M = M_0$ , its uprooted model  $M^+$ , and various rerooted models  $M_i$ , we use the following notation: let  $p_i$  be the probability distribution in model  $M_i$ , in particular  $p_0$  is the distribution for model  $M_0$  which is the original model  $M$ ; let  $p_+$  be the distribution in the uprooted model  $M^+$ .

Using similar reasoning to that used above in the proof of Lemma 2, we use the symmetry of  $M^+$  to show the following results which will be useful in §8.3 for mapping rooted probabilities  $p_i$  to ‘universal’ uprooted probabilities  $p_+$ .

**Lemma 5.** (i) For any distinct  $i, j \in \{0, \dots, n\}$ ,  $p_i(X_j = 1) = p_+(X_i \neq X_j)$ ;  
 (ii) for any distinct  $i, j, k \in \{0, \dots, n\}$ ,  $p_i(X_j \neq X_k) = p_+(X_j \neq X_k)$ .

*Proof.* (i) For distinct  $i, j \in \{0, \dots, n\}$ ,

$$\begin{aligned} p_i(X_j = 1) &= p_+(X_j = 1 | X_i = 0) \\ &= \frac{p_+(X_j = 1, X_i = 0)}{p_+(X_i = 0)} \\ &= 2p_+(X_j = 1, X_i = 0) \\ &= p_+(X_j = 1, X_i = 0) + p_+(X_j = 0, X_i = 1) \quad (\text{symmetry of } M^+) \\ &= p_+(X_i \neq X_j). \end{aligned}$$

(ii) For distinct  $i, j, k \in \{0, \dots, n\}$ ,

$$\begin{aligned} p_i(X_j \neq X_k) &= p_+(X_j \neq X_k | X_i = 0) \\ &= \frac{p_+(X_j = 1, X_k = 0, X_i = 0) + p_+(X_j = 0, X_k = 1, X_i = 0)}{p_+(X_i = 0)} \\ &= 2[p_+(X_j = 1, X_k = 0, X_i = 0) + p_+(X_j = 0, X_k = 1, X_i = 0)] \\ &= p_+(X_j = 1, X_k = 0, X_i = 0) + p_+(X_j = 0, X_k = 1, X_i = 0) \\ &\quad + p_+(X_j = 0, X_k = 1, X_i = 1) + p_+(X_j = 1, X_k = 0, X_i = 1) \quad (\text{symmetry of } M^+) \\ &= p_+(X_j = 1, X_k = 0) + p_+(X_j = 0, X_k = 1) \\ &= p_+(X_j \neq X_k). \end{aligned}$$

□

### 8.1. The marginal polytope and its relaxations LOC and TRI

Given a model  $M$  with  $n$  variables  $\mathcal{V}$  and  $m$  edges  $\mathcal{E}$ , we may consider a vector containing marginal probabilities for all  $n$  single variables and all  $m$  pairs of variables that are directly related.

Specifically, regarding the score (1), for any configuration  $x = (x_1, \dots, x_n)$ , let  $y_{ij} = \mathbb{1}[x_i \neq x_j]$  then collect the  $x$  and  $y$  terms together into a vector  $z = (x_1, \dots, x_n, \dots, y_{ij}, \dots) \in \{0, 1\}^{n+m}$ . Similarly collect together the potential parameters into a vector  $w = (\theta_1, \dots, \theta_n, \dots, -\frac{1}{2}W_{ij}, \dots) \in \mathbb{R}^{n+m}$ . Now the score of a configuration  $x$  may be written as  $w \cdot z(x)$ , and MAP inference may be framed as an integer linear program to find  $z^* \in \arg \max_{z: x \in \{0, 1\}^n} w \cdot z$ .

The convex hull of the  $2^n$  possible integer solutions in  $[0, 1]^{n+m}$  is the *marginal polytope*  $\mathbb{M}$  for our choice of singleton and edge terms in (1). Regarding the convex coefficients as a probability distribution  $p_0$  over all possible states, the marginal polytope may be considered the space of all singleton and pairwise mean marginals that are consistent with some global distribution  $p_0$  over the  $2^n$  states, that is

$$\mathbb{M} = \{\mu = (\mu_1, \dots, \mu_n, \dots, \mu_{ij}, \dots) \in [0, 1]^d \text{ s.t. } \exists p_0 : \mu_i = \mathbb{E}_{p_0}(X_i) \forall i, \mu_{ij} = \mathbb{E}_{p_0}(\mathbb{1}[X_i \neq X_j]) \forall (i, j) \in E\}. \quad (4)$$

Note that  $\mu_i = p_0(X_i = 1)$  and  $\mu_{ij} = p_0(X_i \neq X_j)$ .

Since an LP attains an optimum at a vertex of the feasible region, if  $w \cdot \mu$  is maximized over  $\mathbb{M}$  then an exact integer solution is always optimum. However,  $\mathbb{M}$  has exponentially many facets (Deza and Laurent, 1997), hence a simpler, relaxed constraint set is typically employed, yielding an upper bound on the original optimum. This set is often chosen as the *local polytope* LOC, which enforces only pairwise consistency (Wainwright and Jordan, 2008). If an optimum vertex is achieved at an integer solution, then this must be an optimum of the original discrete problem, in which case we say

that the relaxation LP+LOC is *tight*. Sherali and Adams (1990) proposed a series of successively tighter relaxations by enforcing consistency over progressively larger clusters of variables. At order  $r$ , the  $\mathcal{L}_r$  polytope enforces consistency over all clusters of variables of size  $\leq r$ .  $\mathcal{L}_2$  is the local polytope LOC. Next,  $\mathcal{L}_3$  is the triplet-consistent polytope TRI, and so on, with  $\mathcal{L}_n = \mathbb{M} \subseteq \mathcal{L}_{n-1} \subseteq \dots \subseteq \mathcal{L}_3 = \text{TRI} \subseteq \mathcal{L}_2 = \text{LOC}$ .

In order to obtain the explicit constraints for these polytopes, earlier work (Wainwright and Jordan, 2008; Weller et al., 2016) uses a different (but equivalent) minimal reparameterization leading to a different (but equivalent) set of marginals. To link to their notation, let  $\alpha_i = p_0(X_i = 1)$ ,  $\alpha_{ij} = p_0(X_i = 1, X_j = 1)$ ,  $\alpha_{ijk} = p_0(X_i = 1, X_j = 1, X_k = 1)$ . We next present a derivation of the constraints for LOC and TRI following (Weller et al., 2016), see also (Wainwright and Jordan, 2008, Example 8.7).

Examining just one variable, we have  $\alpha_i = \mu_i \in [0, 1] \forall i$ . In order to be consistent with these single variable marginals, the matrix of pairwise marginals for edge  $(i, j)$  takes the form

$$\begin{pmatrix} p_0(X_i = 0, X_j = 0) & p_0(X_i = 0, X_j = 1) \\ p_0(X_i = 1, X_j = 0) & p_0(X_i = 1, X_j = 1) \end{pmatrix} = \begin{pmatrix} 1 + \alpha_{ij} - \alpha_i - \alpha_j & \alpha_j - \alpha_{ij} \\ \alpha_i - \alpha_{ij} & \alpha_{ij} \end{pmatrix}. \quad (5)$$

The LOC constraints are exactly those that ensure that all 4 terms are  $\geq 0$ , which leads to

$$\text{LOC constraints for edge } (i, j) : \quad \max(0, \alpha_i + \alpha_j - 1) \leq \alpha_{ij} \leq \min(\alpha_i, \alpha_j). \quad (6)$$

These constraints may be reformulated in terms of our  $\mu_{ij}$  marginals by using  $\mu_i = \alpha_i$ , and observing from (5) that

$$\mu_{ij} = \alpha_i + \alpha_j - 2\alpha_{ij} \quad \Leftrightarrow \quad \alpha_{ij} = \frac{1}{2}(\mu_i + \mu_j - \mu_{ij}).^5 \quad (7)$$

To obtain the constraints for TRI, we use a ‘lift-and-project’ approach by ‘lifting’ to distributions over three variables, deriving conditions, then projecting these back down to the one and two variable marginals that we are using. We must ensure that the distribution over every triplet of variables  $X_i, X_j, X_k$  is valid and consistent with all edge and singleton marginals. Given  $\alpha_i, \alpha_j, \alpha_k, \alpha_{ij}, \alpha_{ik}, \alpha_{jk}$  and using  $\alpha_{ijk} = p_0(X_i = 1, X_j = 1, X_k = 1)$  as defined above, we have:

With  $k = 0$ ,

$$\begin{pmatrix} p_0(X_i = 0, X_j = 0) & p_0(X_i = 0, X_j = 1) \\ p_0(X_i = 1, X_j = 0) & p_0(X_i = 1, X_j = 1) \end{pmatrix} = \begin{pmatrix} 1 - \alpha_i - \alpha_j - \alpha_k + \alpha_{ij} + \alpha_{ik} + \alpha_{jk} - \alpha_{ijk} & \alpha_j + \alpha_{ijk} - \alpha_{ij} - \alpha_{jk} \\ \alpha_i + \alpha_{ijk} - \alpha_{ij} - \alpha_{ik} & \alpha_{ij} - \alpha_{ijk} \end{pmatrix}$$

With  $k = 1$ ,

$$\begin{pmatrix} p_0(X_i = 0, X_j = 0) & p_0(X_i = 0, X_j = 1) \\ p_0(X_i = 1, X_j = 0) & p_0(X_i = 1, X_j = 1) \end{pmatrix} = \begin{pmatrix} \alpha_k + \alpha_{ijk} - \alpha_{ik} - \alpha_{jk} & \alpha_{jk} - \alpha_{ijk} \\ \alpha_{ik} - \alpha_{ijk} & \alpha_{ijk} \end{pmatrix}.$$

As previously for LOC, we have the constraints that all terms are  $\geq 0$ . By combining inequalities, we may project back down by eliminating  $\alpha_{ijk}$ . For example, if we combine the condition that the top right element of the matrix for  $k = 0$  is  $\geq 0$  with the similar condition for the bottom right element of the same matrix, we obtain  $\alpha_j - \alpha_{jk} \geq 0$  which is one of the LOC constraints for edge  $(j, k)$ , see (6). Working through the various combinations yields all the previous LOC constraints for the edges  $(i, j)$ ,  $(i, k)$  and  $(j, k)$ , and in addition we obtain the following four new *triplet constraints*, which are called *cycle inequalities* in (Wainwright and Jordan, 2008, Example 8.7).

$$\begin{aligned} \text{TRI constraints in terms of } \alpha \text{ marginals for triplet of distinct } i, j, k \in \{1, \dots, n\} : \quad & \alpha_i + \alpha_{jk} \geq \alpha_{ij} + \alpha_{ik} \\ & \alpha_j + \alpha_{ik} \geq \alpha_{ij} + \alpha_{jk} \\ & \alpha_k + \alpha_{ij} \geq \alpha_{ik} + \alpha_{jk} \\ & \alpha_{ij} + \alpha_{ik} + \alpha_{jk} \geq \alpha_i + \alpha_j + \alpha_k - 1. \end{aligned} \quad (8)$$

<sup>5</sup>This equivalence is essentially the *covariance mapping* described in (Deza and Laurent, 1997, §5.2).

If we use (7) to rewrite these TRI constraints (8) in terms of  $\mu$  marginals, then they take the following appealing form.

$$\begin{aligned}
 \text{TRI constraints in terms of } \mu \text{ marginals for triplet of distinct } i, j, k \in \{1, \dots, n\} : \quad & \mu_{jk} \leq \mu_{ij} + \mu_{ik} \\
 & \mu_{ik} \leq \mu_{ij} + \mu_{jk} \\
 & \mu_{ij} \leq \mu_{ik} + \mu_{jk} \\
 & \mu_{ij} + \mu_{ik} + \mu_{jk} \leq 2. \tag{9}
 \end{aligned}$$

Notice that (9) considers only terms of the form  $\mu_{ij}$ . Since  $\mu_{ij}$  is the probability that  $X_i$  and  $X_j$  take different values, a simple way to see that the inequalities of (11) are valid is to observe that they clearly hold for any integer settings of  $X_i, X_j, X_k \in \{0, 1\}^3$ , and hence they must hold for any valid probability distribution over the 8 possible settings of these three variables (since this yields a convex combination).

## 8.2. The cut polytope and its relaxations RMET and MET

As in §3: given a model  $M$  with variables  $\{X_1, \dots, X_n\}$  on graph  $G(\mathcal{V}, \mathcal{E})$  with vertices  $\mathcal{V} = \{1, \dots, n\}$  and edges  $\mathcal{E}$ , its uprooted model  $M^+$  has variables  $\{X_0, \dots, X_n\}$  on graph  $G'(\mathcal{V}', \mathcal{E}')$  with vertices  $\mathcal{V}' = \{0, 1, \dots, n\}$  and edges  $\mathcal{E}' = \mathcal{E} \cup \mathcal{F}$ , where  $\mathcal{F} = \{(0, i) : \theta_i \neq 0\}$ . An uprooted model  $M^+$  is completely symmetric. The score (2) considers only edges and examines only whether the end variables of each edge take the same value.

Given a subset  $S \subseteq \mathcal{V}' = \{0, 1, \dots, n\}$ , let  $\delta(S) \in \{0, 1\}^{|\mathcal{E}'|}$  be the *cut vector* of edges of  $\mathcal{E}'$  which run between the vertex partitions  $S$  and  $\mathcal{V}' \setminus S$ , defined by  $\delta(S)_{ij} = 1$  iff  $i$  and  $j$  are in different partitions.

The *cut polytope* (Barahona and Mahjoub, 1986) of  $G'$  is the convex hull of all such cut vectors, that is  $\text{CUT} = \text{conv} \{ \delta(S) : S \subseteq \mathcal{V}' \}$ . Although there are  $2^{n+1}$  choices of  $S$ , CUT has  $2^n$  vertices since by definition  $\delta(S) = \delta(\mathcal{V}' \setminus S)$ . In fact, there is a simple linear bijection between CUT and the marginal polytope  $\mathbb{M}$  of  $M$ .

Given  $d \in \text{CUT}$  with entries  $d(i, j)$  for each edge  $(i, j) \in \mathcal{E}'$ ,  $d$  maps to  $\mu \in \mathbb{M}$  where  $\mu_j = d(0, j)$  for  $j \in \mathcal{V}$ , and  $\mu_{ij} = d(i, j)$  for  $(i, j) \in \mathcal{E}$ . To see this,  $d(i, j)$  may be interpreted as the marginal probability that  $i, j \in \mathcal{V}'$  lie in different partitions.

As an aside, note that the marginal polytope of  $M^+$ , which we call  $\mathbb{M}^+$ , is closely related, but different, to CUT.  $\mathbb{M}^+$  has  $n + 1$  additional dimensions for the singleton marginal dimensions of its  $n + 1$  variables, though given the symmetry of  $M^+$ , these are all 1/2.

MAP inference for the model  $M$  on  $G$  is equivalent to the weighted max cut problem for  $G'$ :

$$\max_{\mu \in \mathbb{M}} w \cdot \mu = \max_{e \in \text{CUT}} w' \cdot d, \quad w'_{ij} = \begin{cases} \theta_j & i = 0 \\ -\frac{1}{2}W_{ij} & (i, j) \in E. \end{cases} \tag{10}$$

The bijection between  $\mathbb{M}$  and CUT may also be used to map the LOC and TRI relaxations of  $\mathbb{M}$  to corresponding relaxations of CUT in  $[0, 1]^{|\mathcal{E}'|}$ , called the *rooted semimetric polytope* RMET and the *semimetric polytope* MET, respectively. The constraints for the MET polytope (which corresponds to TRI) take the following form, sometimes described as unrooted triangle inequalities (Deza and Laurent, 1997, §27.1):

$$\begin{aligned}
 \text{MET constraints } \forall \text{ distinct } i, j, k \in \mathcal{V}' = \{0, 1, \dots, n\} : \quad & d(i, j) - d(i, k) - d(j, k) \leq 0 \\
 & d(i, j) + d(i, k) + d(j, k) \leq 2. \tag{11}
 \end{aligned}$$

Note that the MET constraints (11) restricted to triplets  $i, j, k \in \mathcal{V} = \{1, \dots, n\}$  are identical to the TRI constraints for  $\mu$  marginals in (9). Both enforce triplet consistency on the marginal probabilities of edges having end vertices which are different.

Remarkably, the constraints on  $d$  for RMET, the *rooted* triangle inequalities, which are equivalent to the LOC constraints on  $\mu$  for LOC (6), are exactly just those of (11) for which one of  $i, j, k$  is 0, the vertex that was added to  $G$  to yield  $G'$ . Hence, RMET may be regarded as MET *rooted* at 0. Correspondingly, we may consider TRI to be a version of LOC that is *universally rooted*.

To see this, we shall consider the LOC constraints for edge  $(i, j) \in \mathcal{E}$  (6), and show that they are exactly the MET constraints (11) applied to triplet  $(0, i, j)$  in  $\mathcal{V}'$ . Consider the triangle  $0ij$  of  $G'$  shown in Figure 4.

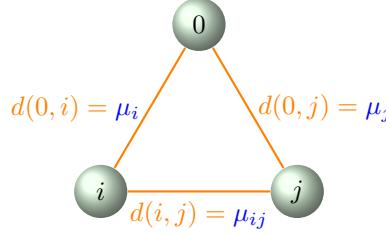


Figure 4. Illustration of edge marginals for the MET polytope, shown in orange, and their values in terms of  $\mu$  marginals for the model  $M$ , shown in blue.

Recall that  $\mu_i = \alpha_i$ ,  $\mu_j = \alpha_j$ , and from (7) that  $\mu_{ij} = \alpha_i + \alpha_j - 2\alpha_{ij}$ . Hence, the MET constraints (11) with  $k = 0$  become:

$$\begin{aligned} d(i, j) - d(i, k) - d(j, k) \leq 0 &\Leftrightarrow \alpha_i + \alpha_j - 2\alpha_{ij} - \alpha_i - \alpha_j \leq 0 &\Leftrightarrow \alpha_{ij} \geq 0 \\ d(i, k) - d(i, j) - d(j, k) \leq 0 &\Leftrightarrow \alpha_i - \alpha_i - \alpha_j + 2\alpha_{ij} - \alpha_j \leq 0 &\Leftrightarrow \alpha_{ij} \leq \alpha_j \\ d(j, k) - d(i, j) - d(i, k) \leq 0 &\Leftrightarrow \alpha_j - \alpha_i - \alpha_j + 2\alpha_{ij} - \alpha_i \leq 0 &\Leftrightarrow \alpha_{ij} \leq \alpha_i \\ d(i, j) + d(i, k) + d(j, k) \leq 2 &\Leftrightarrow \alpha_i + \alpha_j - 2\alpha_{ij} + \alpha_i + \alpha_j \leq 2 &\Leftrightarrow \alpha_{ij} \geq \alpha_i + \alpha_j - 1, \end{aligned}$$

which exactly match the LOC constraints (6), as required.

### 8.3. New results

With the background in §8.1-8.2, we are ready to prove our new results.

**Notation.** Let  $\mu^i, w^i$  be the  $\mu, w$  vectors corresponding to rerootings at  $X_i$ . In particular,  $\mu^0, w^0$  are the  $\mu, w$  vectors for the original model  $M = M_0$ .

**Theorem 3. (TRI is ‘universally rooted’)** LP+TRI yields the same optimum score for  $M$  as for any rerooting  $M_i$ ; hence LP+TRI is either tight for all rerootings or for none.

*Proof.* First, note that the MAP score for  $M$  is the same as that for any rerooting  $M_i$ . One way to see this follows the observations in §3-4: each configuration  $x$  of  $M$  maps to exactly 2 configurations of  $M^+$ :  $y_0 = (0, x)$  and  $y_1 = \bar{y}_0 = (1, \bar{x})$ , with the potentials of  $M^+$  set so that  $\text{score}(x) = \text{score}(y_0) = \text{score}(y_1)$ . Hence, in particular, a MAP configuration for  $M$  maps to two MAP configurations for  $M^+$  with the same score, and exactly one of these will be in any rerooting  $M_i$  as a MAP configuration for that rerooted model with the same score.

It remains to show that  $\max_{\text{TRI}(M_i)} w^i \cdot \mu^i$  is the same for any rerooting of a model  $M$ . We shall use a similar idea, converting the problem for  $M$  into a problem over the graph  $G'$  of the uprooted model, in such a way that this problem over  $G'$  is the same for all rerootings  $M_i$ . In fact, we shall show a score-preserving linear bijection between  $\text{TRI}(M_i)$  and MET, where we must still show that this is the same no matter which rerooting  $M_i$  is used.

In §8.2, we gave a simple linear bijection between  $\mathbb{M}$  and CUT, which naturally extends to a linear bijection between  $\text{TRI}(M)$  and  $\text{MET}(M)$ . Further it is clear that this is score preserving if we use  $w'$  from (10). That is, we have for any  $\mu \in \text{TRI}(M)$ , a linear bijection between  $\mu$  and  $d \in \text{MET}(M)$  s.t.  $w \cdot \mu = w' \cdot d$ . If these are maximized over their respective (equivalent) polytopes, then we obtain the same maximum.

It remains to show that for all rerootings,  $\text{MET}(M) = \text{MET}(M_i)$  and that  $w^i$  maps to the same vector  $w'$  for each MET.  $\text{MET}(M) = \text{MET}(M_i)$  follows directly from Lemma 5. Each  $w^i$  maps to the same vector  $w'$  by construction, see (2).  $\square$

The next result follows as a simple application of Theorem 3 to the earlier result of Weller et al. (2016).

**Theorem 4.** LP+TRI is tight for (any rerooting of) a model  $M$  whose uprooted model  $M^+$  is 2-almost balanced.

*Proof.* First, if  $M^+$  is 2-almost balanced with special variables  $X_i$  and  $X_j$ , then if we root at either  $X_i$  or  $X_j$ , we obtain an almost balanced model (that is,  $M_i$  or  $M_j$ ) on which LP+TRI is tight by the result of Weller et al. (2016). Now if LP+TRI is tight for  $M_i$ , then by Theorem 3, LP+TRI is tight for any rerooting of  $M_i$ , including  $M$ .  $\square$

Following our result, [Weller \(2016\)](#) demonstrated a still stronger result: LP+TRI is tight for any model  $M$  whose uprooted model  $M^+$  does not contain an *odd- $K_5$*  as a *signed minor*. An odd- $K_5$  is the complete graph on 5 variables where all edges are repulsive. Since an odd- $K_5$  is clearly not 2-almost balanced (if any 2 variables are removed, the remaining model is a frustrated triangle), all 2-almost balanced models are a subset of those that do not contain an odd- $K_5$  as a signed minor. Further, the condition of [Weller \(2016\)](#) was shown to be both sufficient and necessary for tightness for models with all potentials that respect the edge signs of the uprooted model. For details, see ([Weller, 2016](#)).

## 9. Details of experimental methods, and additional results

For all inference methods, we used the open source libDAI library ([Mooij, 2010](#)) and averaged over 100 random models. We show results first for smaller models (complete graph on 10 variables and  $5 \times 5$  grids), and then in Figure 11 for Bethe for larger models (complete graph on 15 variables and  $9 \times 9$  grids). Wherever possible, we were consistent with the approaches of [Weller and Domke \(2016\)](#). We experienced difficulty with mean field (MF), since a randomly initialized run could return a very suboptimal solution. Hence, each time we used 100 random initializations and took the solution with highest estimate of the partition function (the most accurate since MF always provides a lower bound). Still, we experienced some convergence difficulties and advise caution in interpreting our MF results.

For MF,

```
MF [tol=1e-7, maxiter=10000, damping=0.0, init=RANDOM, updates=NAIVE]
```

For Bethe,

```
HAK [doubleloop=1, clusters=BETHE, init=UNIFORM, tol=1e-7, maxiter=10000]
```

This is guaranteed to converge to a stationary point of the Bethe free energy (whereas BP may not converge).

For TRW,

```
TRWBP [updates=SEQFIX, tol=1e-7, maxiter=10000, logdomain=0, nrtrees=1000, ...
damping=0.25, init=UNIFORM]
```

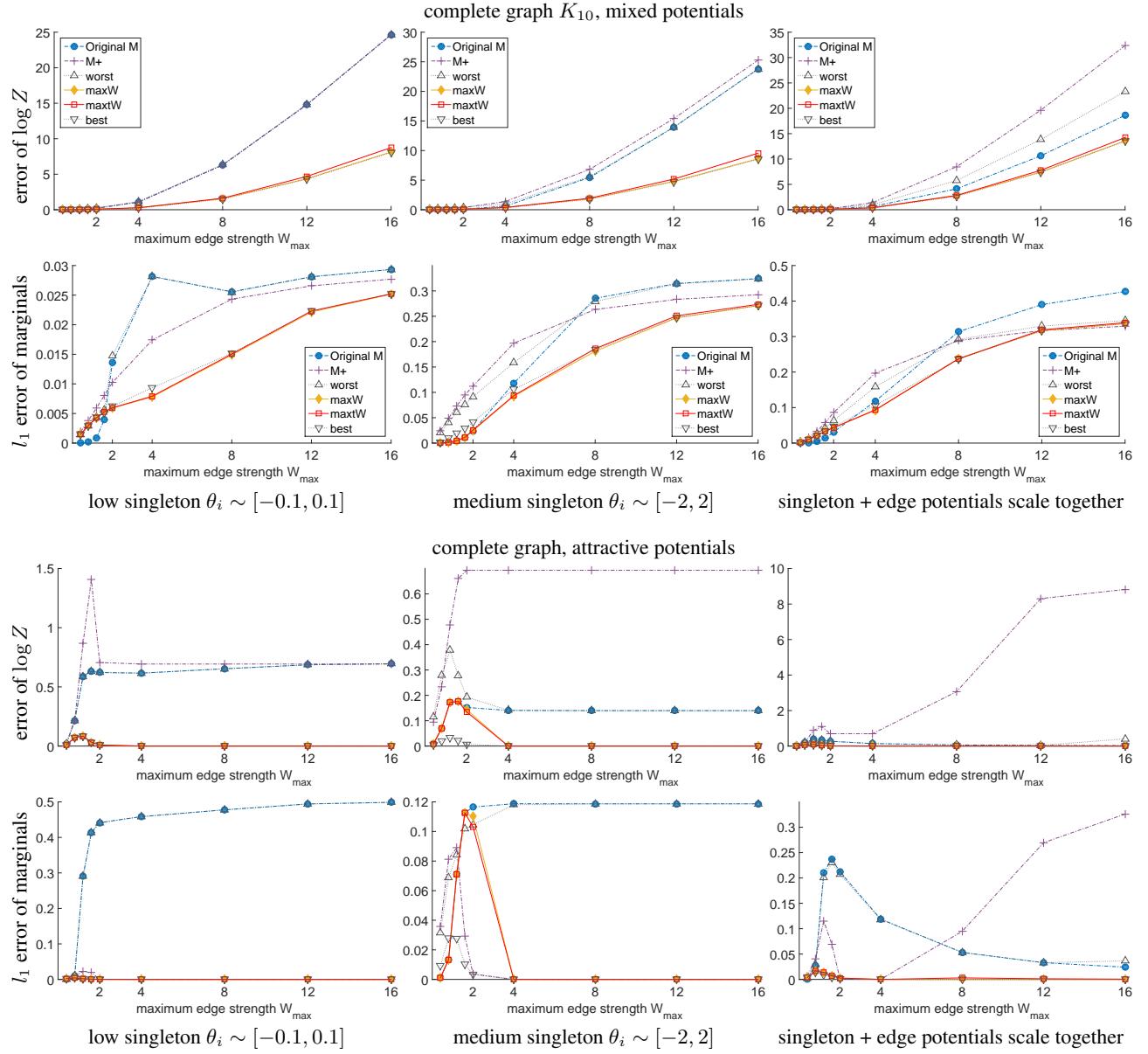


Figure 5. Average error plots over 100 runs for the Bethe approximation, complete graph with 10 variables

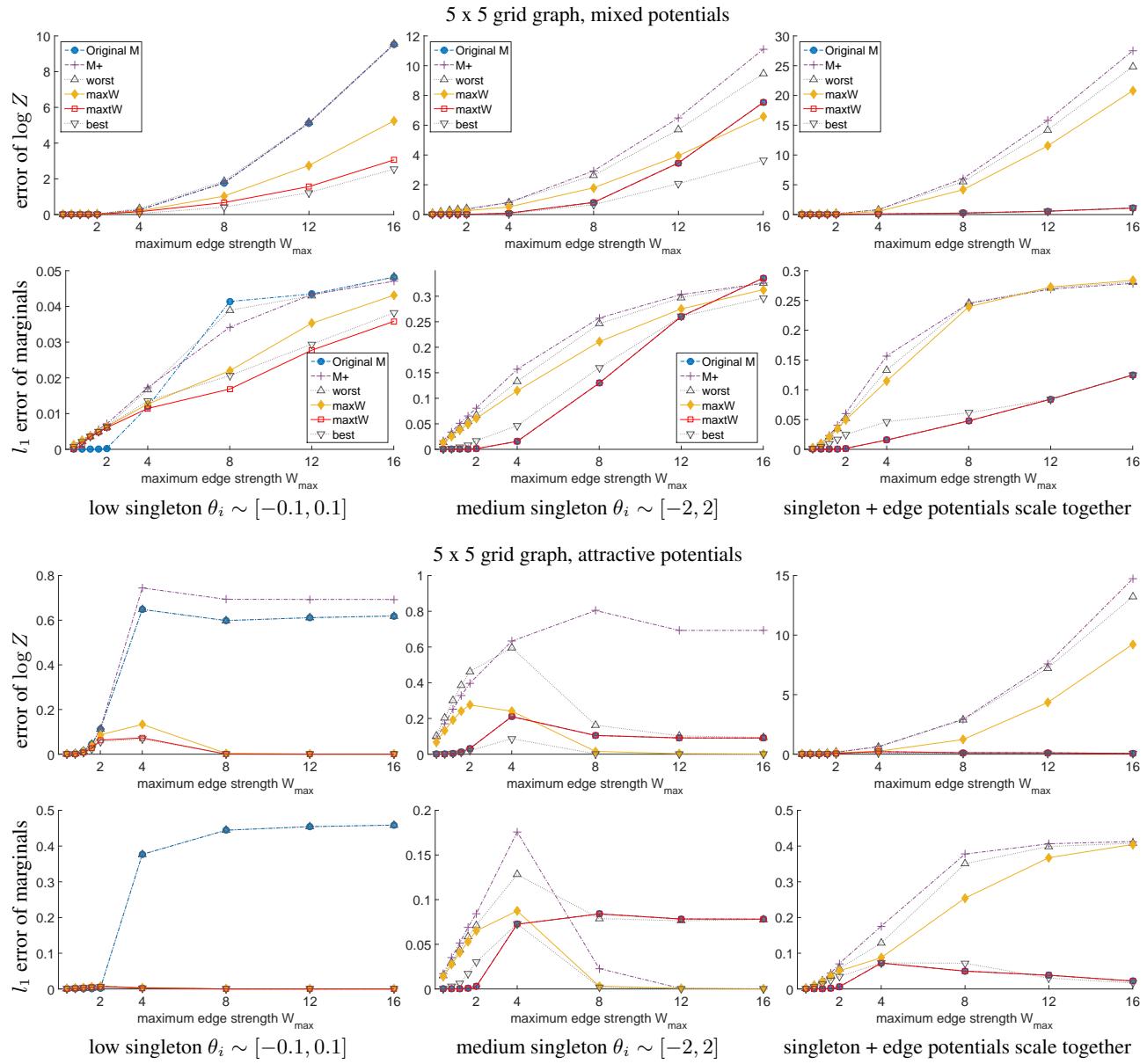


Figure 6. Average error plots over 100 runs for the Bethe approximation, 5 x 5 grid graph

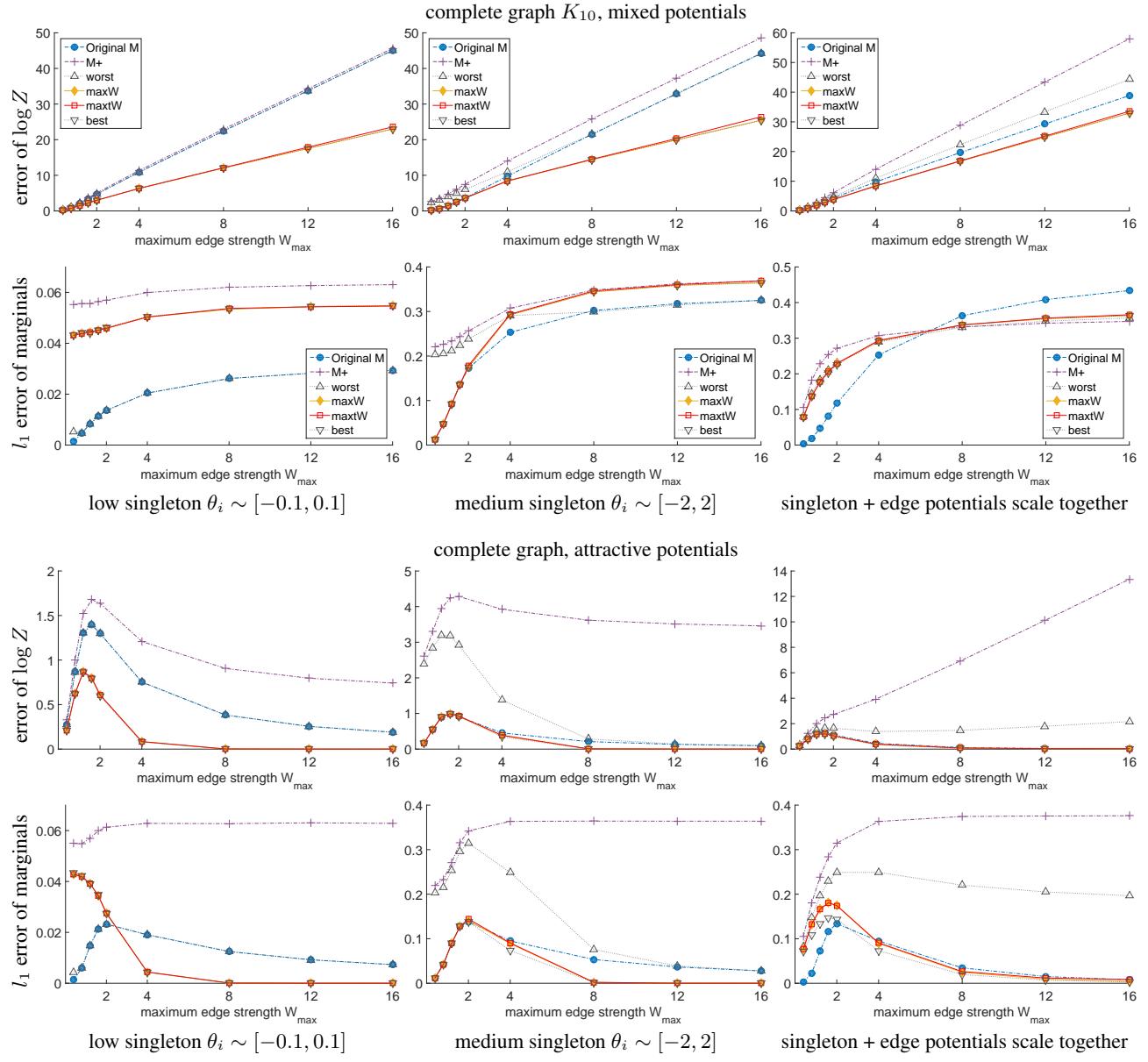


Figure 7. Average error plots over 100 runs for the TRW approximation, complete graph with 10 variables  
**Note the very low scale for  $l_1$  error of marginals for low singleton potentials.**

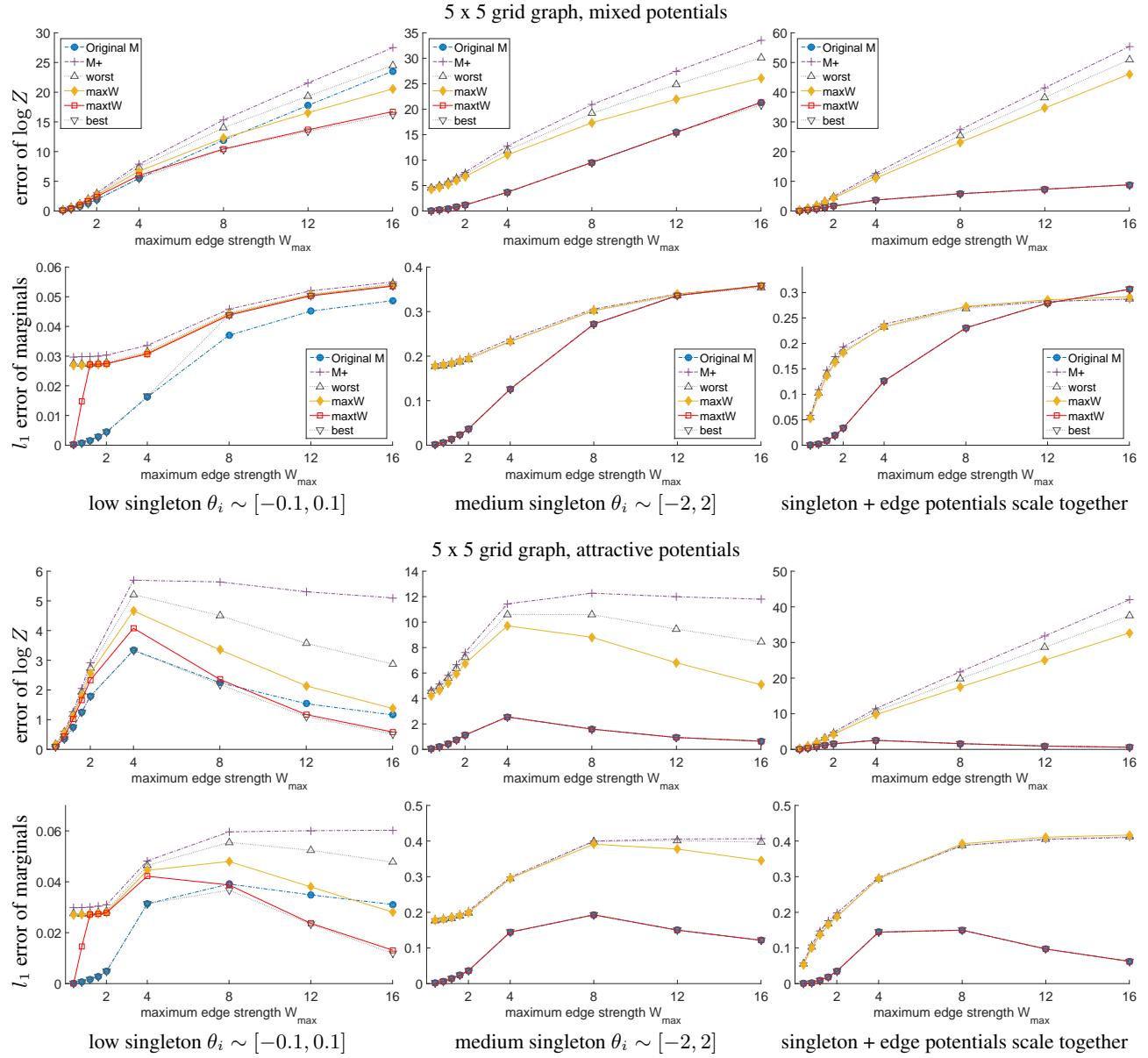


Figure 8. Average error plots over 100 runs for the TRW approximation, 5 x 5 grid graph  
**Note the very low scale for  $l_1$  error of marginals for low singleton potentials.**

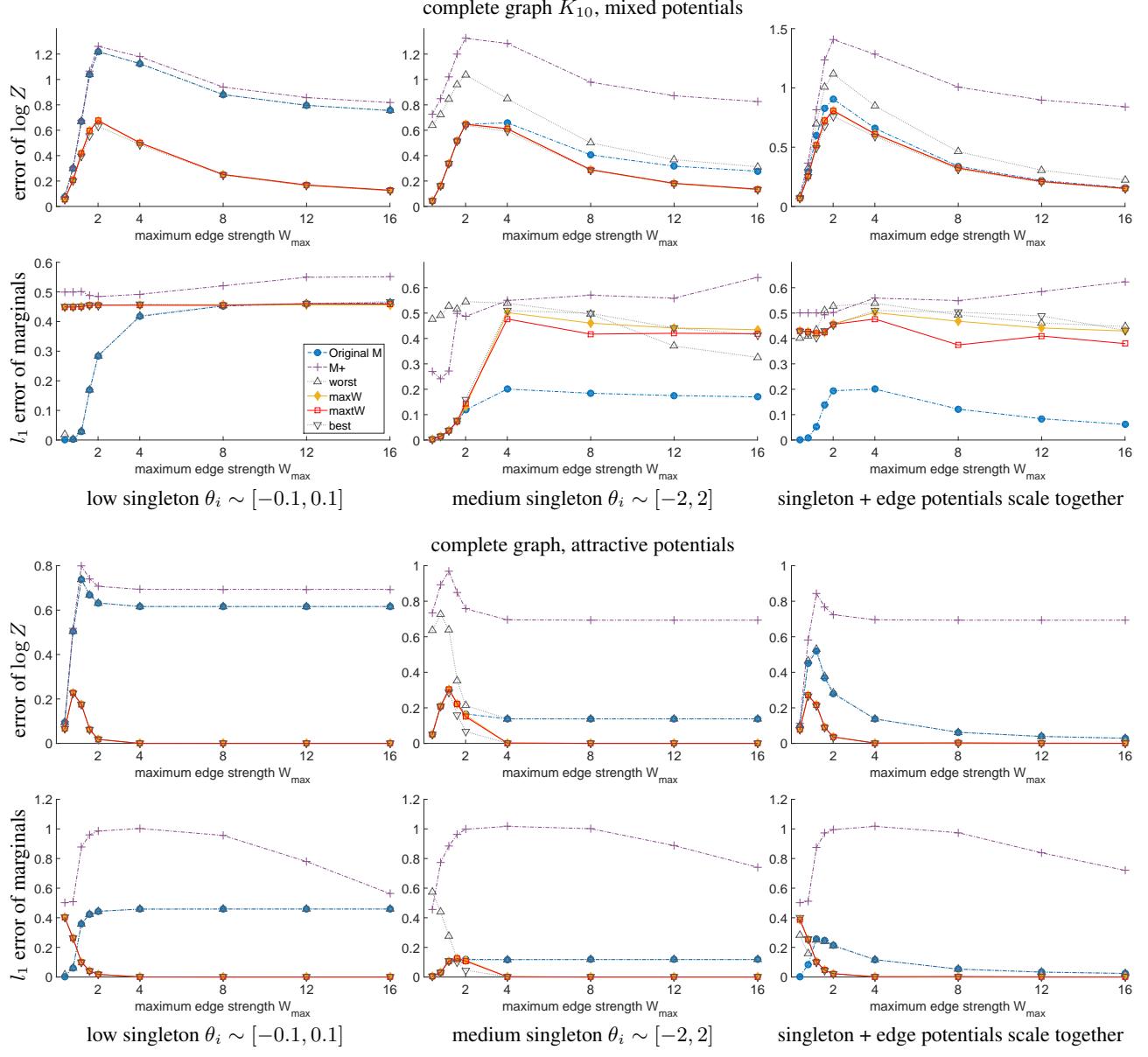


Figure 9. Average error plots over 100 runs for the MF approximation, complete graph with 10 variables

**Results for the error of marginals for the complete graph look interesting and warrant further investigation, though we suspect these may be due to problems with our MF algorithm implementation.**

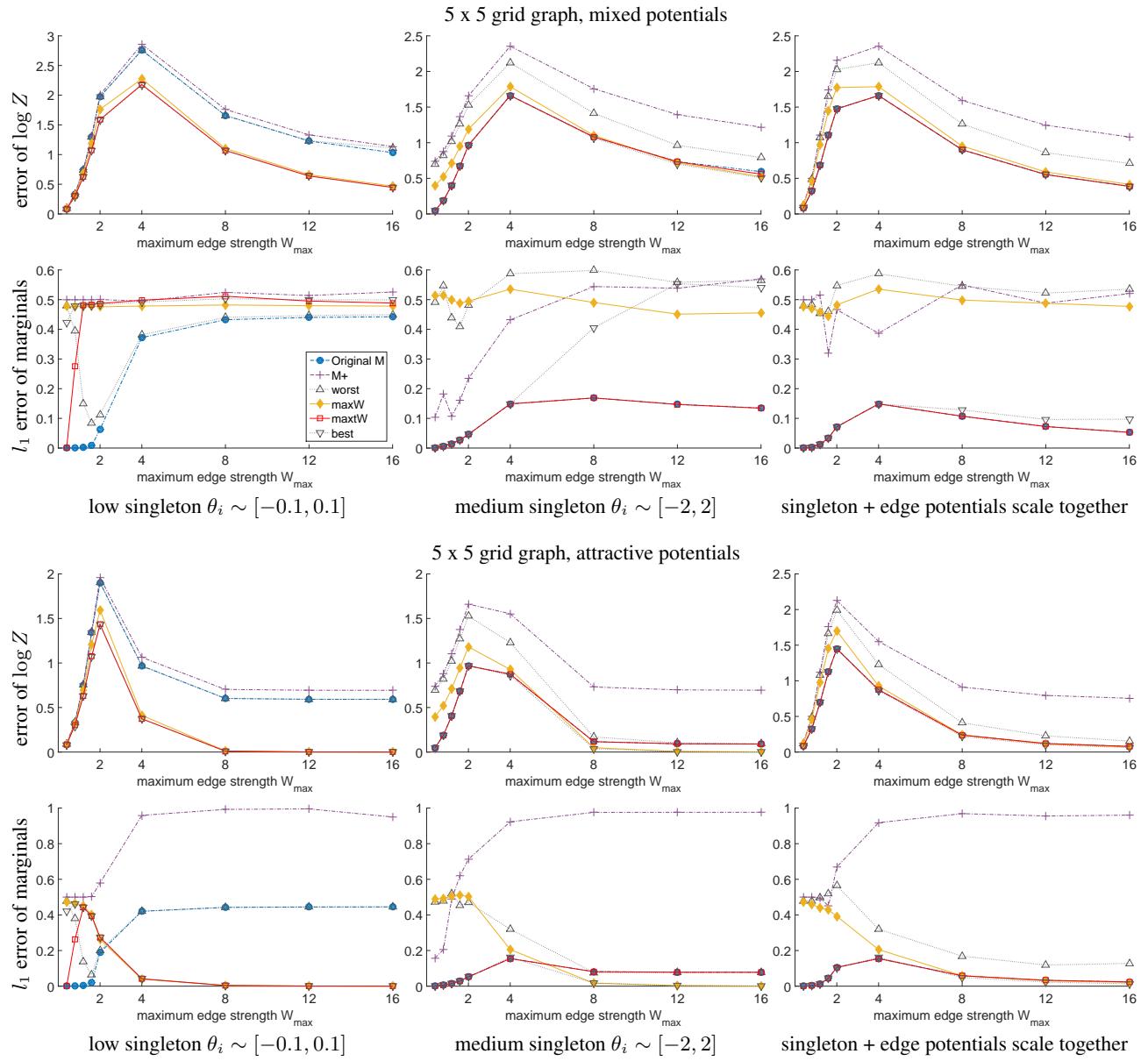


Figure 10. Average error plots over 100 runs for the MF approximation, 5 x 5 grid graph

## Bethe results for larger models, mixed potentials

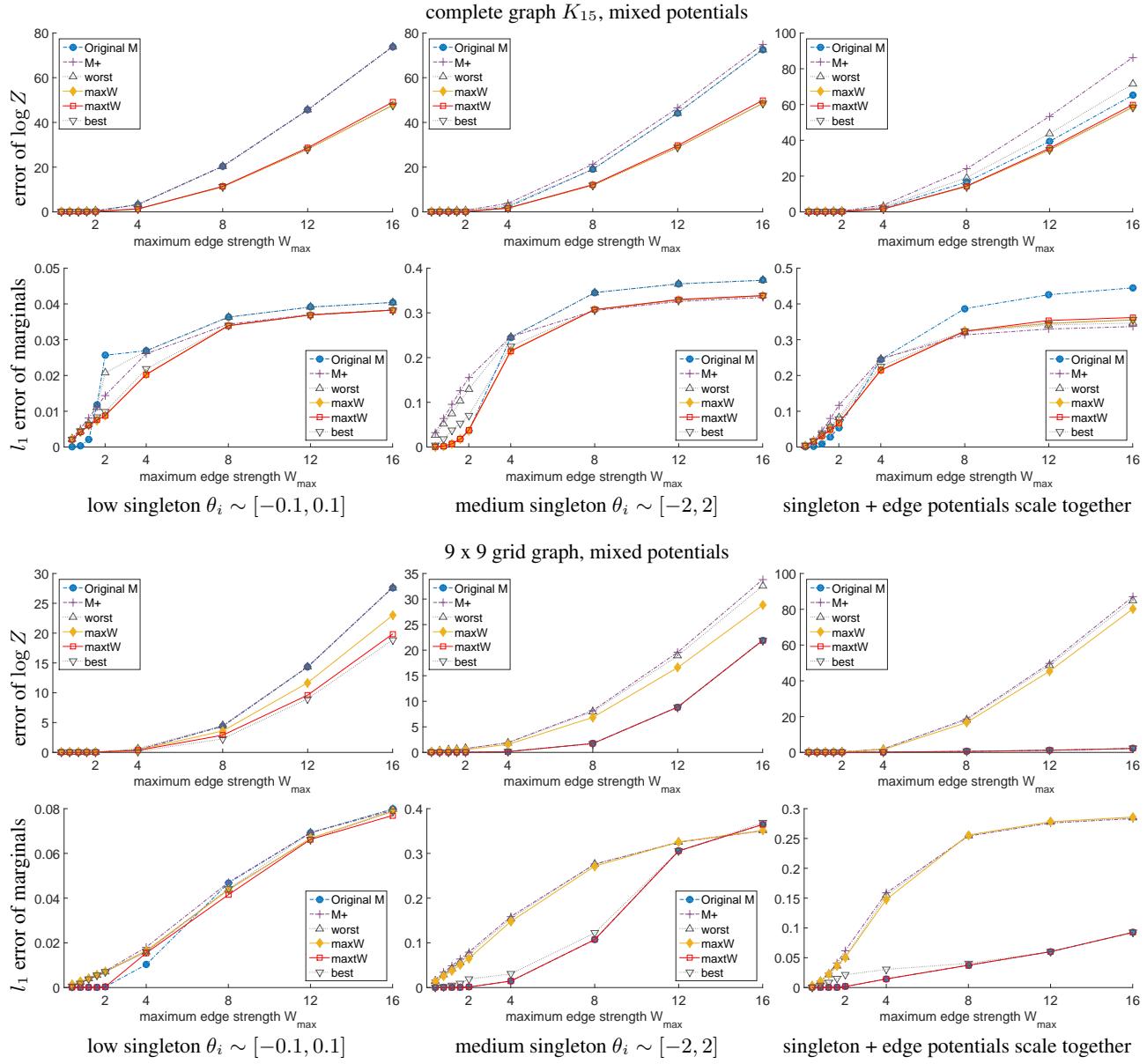


Figure 11. Average error plots over 100 runs for the Bethe approximation, mixed potentials