Lecture 6: Graphical Models: Learning 4F13: Machine Learning

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Learning parameters



Assume each variable x_i is discrete and can take on K_i values.

The parameters of this model can be represented as 4 tables: θ_1 has K_1 entries, θ_2 has $K_1 \times K_2$ entries, etc.

These are called **conditional probability tables** (CPTs) with the following semantics:

 $P(x_1)P(x_2|x_1)P(x_3|x_1)P(x_4|x_2)$

$$P(x_1 = k) = \theta_{1,k}$$
 $P(x_2 = k' | x_1 = k) = \theta_{2,k,k'}$

If node *i* has *M* parents, θ_i can be represented either as an *M* + 1 dimensional table, or as a 2-dimensional table with $\left(\prod_{i \in pa(i)} K_i\right) \times K_i$ entries by collapsing all the states of the parents of node *i*. Note that $\sum_{k'} \theta_{i,k,k'} = 1$.

Assume a data set $\mathcal{D} = {\mathbf{x}^{(n)}}_{n=1}^{N}$. How do we learn θ from \mathcal{D} ?

Learning parameters

Assume a data set $\mathcal{D} = {\mathbf{x}^{(n)}}_{n=1}^N$. How do we learn $\boldsymbol{\theta}$ from \mathcal{D} ?

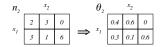
 $P(\mathbf{x}|\theta) = P(x_1|\theta_1)P(x_2|x_1,\theta_2)P(x_3|x_1,\theta_3)P(x_4|x_2,\theta_4)$ Likelihood: $P(\mathcal{D}|\boldsymbol{\theta}) = \prod^{N} P(\mathbf{x}^{(n)}|\boldsymbol{\theta})$

Log Likelihood:

$$\log P(\mathcal{D}|\boldsymbol{\theta}) = \sum_{n=1}^{N} \sum_{i=1}^{n=1} \log P(x_i^{(n)}|x_{\mathrm{pa}(i)}^{(n)}, \boldsymbol{\theta}_i)$$

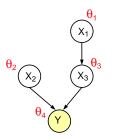
This decomposes into sum of functions of θ_i . Each θ_i can be optimized separately: $\hat{\theta}_{i,k,k'} = \frac{n_{i,k,k'}}{\sum_{k,l'} n_{i,k,k''}}$

where $n_{i,k,k'}$ is the number of times in \mathcal{D} where $x_i = k'$ and $x_{pa(i)} = k$.



ML solution: Simply calculate frequencies!





Assume a model parameterised by θ with observable variables *Y* and hidden variables *X*

Goal: maximize parameter log likelihood given observed data.

$$\mathcal{L}(\theta) = \log p(Y|\theta) = \log \sum_{X} p(Y, X|\theta)$$

Goal: maximise parameter log likelihood given observables.

$$\mathcal{L}(\theta) = \log p(Y|\theta) = \log \sum_{X} p(Y, X|\theta)$$

The EM algorithm (intuition):

Iterate between applying the following two steps:

- The E step: fill-in the hidden/missing variables
- The M step: apply complete data learning to filled-in data.

Goal: maximise parameter log likelihood given observables.

$$\mathcal{L}(\theta) = \log p(Y|\theta) = \log \sum_{X} p(Y, X|\theta)$$

The EM algorithm (derivation):

$$\mathcal{L}(\theta) = \log \sum_{X} q(X) \frac{p(Y, X|\theta)}{q(X)} \ge \sum_{X} q(X) \log \frac{p(Y, X|\theta)}{q(X)} = \mathcal{F}(q(X), \theta)$$

- The E step: maximize $\mathcal{F}(q(X), \theta^{[t]})$ wrt q(X) holding $\theta^{[t]}$ fixed: $q(X) = P(X|Y, \theta^{[t]})$
- The M step: maximize $\mathcal{F}(q(X), \theta)$ wrt θ holding q(X) fixed:

$$\theta^{[t+1]} \leftarrow \operatorname{argmax}_{\theta} \sum_{X} q(X) \log p(Y, X|\theta)$$

The E-step requires solving the *inference* problem, finding the distribution over the hidden variables $p(X|Y, \theta^{[t]})$ given the current model parameters. This can be done using belief propagation or the junction tree algorithm.

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Lecture 6: Graphical Models: Learning

ML Learning with Complete Data (No Hidden Variables)

Log likelihood decomposes into sum of functions of θ_i . Each θ_i can be optimized separately:

$$\hat{\theta}_{ijk} \leftarrow \frac{n_{ijk}}{\sum_{k'} n_{ijk'}}$$

where n_{ijk} is the number of times in \mathcal{D} where $x_i = k$ and $x_{pa(i)} = j$. Maximum likelihood solution: Simply calculate frequencies!

ML Learning with Incomplete Data (i.e. with Hidden Variables)

Iterative EM algorithm

E step: compute expected counts given previous settings of parameters $E[n_{ijk}|\mathcal{D}, \boldsymbol{\theta}^{[t]}]$. **M step:** re-estimate parameters using these expected counts

$$\boldsymbol{\theta}_{ijk}^{[t+1]} \leftarrow \frac{E[\boldsymbol{n}_{ijk} | \mathcal{D}, \boldsymbol{\theta}^{[t]}]}{\sum_{k'} E[\boldsymbol{n}_{ijk'} | \mathcal{D}, \boldsymbol{\theta}^{[t]}]}$$

Bayesian parameter learning with no hidden variables

Let
$$n_{ijk}$$
 be the number of times $(x_i^{(n)} = k \text{ and } x_{pa(i)}^{(n)} = j)$ in \mathcal{D} .

For each *i* and *j*, θ_{ij} is a probability vector of length $K_i \times 1$. Since x_i is a discrete variable with probabilities given by $\theta_{i,j}$, the likelihood is:

$$P(\mathcal{D}|\boldsymbol{\theta}) = \prod_{n} \prod_{i} P(x_{i}^{(n)}|x_{\mathrm{pa}(i)}^{(n)}, \boldsymbol{\theta}) = \prod_{i} \prod_{j} \prod_{k} \theta_{ijk}^{n_{ijk}}$$

If we choose a prior on θ of the form:

$$P(\mathbf{\Theta}) = c \prod_{i} \prod_{j} \prod_{k} \Theta_{ijk}^{\alpha_{ijk}-1}$$

where *c* is a normalization constant, and $\sum_k \theta_{ijk} = 1 \ \forall i, j$, then the posterior distribution also has the same form:

$$P(\boldsymbol{\theta}|\mathcal{D}) = c' \prod_{i} \prod_{j} \prod_{k} \theta_{ijk}^{\tilde{\alpha}_{ijk}-1}$$

where $\tilde{\alpha}_{ijk} = \alpha_{ijk} + n_{ijk}$.

This distribution is called the Dirichlet distribution.

Dirichlet Distribution

The Dirichlet distribution is a distribution over the *K*-dim probability simplex. Let θ be a *K*-dimensional vector s.t. $\forall j : \theta_j \ge 0$ and $\sum_{i=1}^{K} \theta_i = 1$

$$P(\boldsymbol{\theta}|\boldsymbol{\alpha}) = \operatorname{Dir}(\boldsymbol{\alpha}_1, \dots, \boldsymbol{\alpha}_K) \stackrel{\text{def}}{=} \frac{\Gamma(\sum_j \alpha_j)}{\prod_j \Gamma(\alpha_j)} \prod_{j=1}^K \boldsymbol{\theta}_j^{\alpha_j - 1}$$

where the first term is a normalization constant¹ and $E(\theta_j) = \alpha_j / (\sum_k \alpha_k)$ The Dirichlet is conjugate to the multinomial distribution. Let

 $x|\theta \sim \text{Multinomial}(\cdot|\theta)$

That is, $P(x = j | \theta) = \theta_j$. Then the posterior is also Dirichlet:

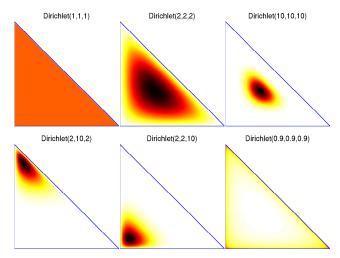
$$P(\boldsymbol{\theta}|\boldsymbol{x}=\boldsymbol{j},\boldsymbol{\alpha}) = \frac{P(\boldsymbol{x}=\boldsymbol{j}|\boldsymbol{\theta})P(\boldsymbol{\theta}|\boldsymbol{\alpha})}{P(\boldsymbol{x}=\boldsymbol{j}|\boldsymbol{\alpha})} = \mathrm{Dir}(\tilde{\boldsymbol{\alpha}})$$

where $\tilde{\alpha}_j = \alpha_j + 1$, and $\forall \ell \neq j : \tilde{\alpha}_\ell = \alpha_\ell$

 ${}^{1}\Gamma(x) = (x-1)\Gamma(x-1) = \int_{0}^{\infty} t^{x-1}e^{-t}dt$. For integer $n, \Gamma(n) = (n-1)!$

Dirichlet Distributions

Examples of Dirichlet distributions over $\theta = (\theta_1, \theta_2, \theta_3)$ which can be plotted in 2D since $\theta_3 = 1 - \theta_1 - \theta_2$:



Example

Assume $\alpha_{ijk} = 1 \ \forall i, j, k$.

This corresponds to a **uniform** prior distribution over parameters θ . This is not a very strong/dogmatic prior, since any parameter setting is assumed a priori possible.

After observed data \mathcal{D} , what are the parameter posterior distributions?

$$P(\theta_{ij} | \mathcal{D}) = \mathrm{Dir}(n_{ij} + 1)$$

This distribution predicts, for future data:

$$P(x_i = k | x_{\text{pa}(i)} = j, \mathcal{D}) = \frac{n_{ijk} + 1}{\sum_{k'} (n_{ijk'} + 1)}$$

Adding 1 to each of the counts is a form of smoothing called "Laplace's Rule".

Bayesian parameter learning with hidden variables

Notation: let \mathcal{D} be the observed data set, *X* be hidden variables, and θ be model parameters. Assume discrete variables and Dirichlet priors on θ

Goal: to infer
$$P(\boldsymbol{\theta}|\mathcal{D}) = \sum_{X} P(X, \boldsymbol{\theta}|\mathcal{D})$$

Problem: since (a)

$$P(\boldsymbol{\theta}|\mathcal{D}) = \sum_{X} P(\boldsymbol{\theta}|X, \mathcal{D}) P(X|\mathcal{D}),$$

and (b) for every way of filling in the missing data, $P(\theta|X, D)$ is a Dirichlet distribution, and (c) there are exponentially many ways of filling in X, it follows that $P(\theta|D)$ is a mixture of Dirichlets with exponentially many terms!

Solutions:

- Find a single best ("Viterbi") completion of *X* (Stolcke and Omohundro, 1993)
- Markov chain Monte Carlo methods
- Variational Bayesian methods (Beal and Ghahramani, 2003)

Summary of parameter learning

	Complete (fully observed) data	Incomplete (hidden /missing) data
ML	calculate frequencies	EM
Bayesian	update Dirichlet distributions	MCMC / Viterbi / VBEM

- For complete data, Bayesian learning is not more costly than ML
- For incomplete data, VBEM \approx EM time complexity
- Other parameter priors are possible but Dirichlet is flexible and intuitive.
- For binary data, other parametrizations include:
 - Sigmoid:

$$P(x_i = 1 | x_{pa(i)}, \theta_i) = 1/(1 + \exp\{-\theta_{i0} - \sum_{j \in pa(i)} \theta_{ij}x_j\})$$

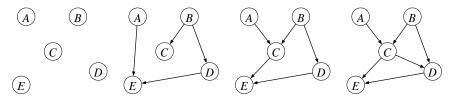
• Noisy-or:

$$P(x_i = 1 | x_{pa(i)}, \theta_i) = 1 - \exp\{-\theta_{i0} - \sum_{j \in pa(i)} \theta_{ij} x_j\}$$

• For non-discrete data, similar ideas but generally harder inference and learning.

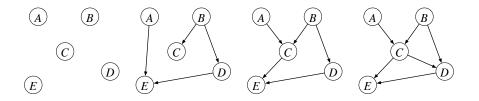
Structure learning

Given a data set of observations of (A, B, C, D, E) can we learn the structure of the graphical model?



Let *m* denote the graph structure = the set of edges.

Structure learning



Constraint-Based Learning: Use statistical tests of marginal and conditional independence. Find the set of DAGs whose d-separation relations match the results of conditional independence tests.

Score-Based Learning: Use a global score such as the BIC score or Bayesian marginal likelihood. Find the structures that maximize this score.

Score-based structure learning for complete data

Consider a graphical model with structure m, discrete observed data D, and parameters θ . Assume Dirichlet priors.

The Bayesian marginal likelihood score is easy to compute:

$$\operatorname{score}(m) = \log P(\mathcal{D}|m) = \log \int P(\mathcal{D}|\theta, m) P(\theta|m) d\theta$$
$$\operatorname{score}(m) = \sum_{i} \sum_{j} \left[\log \Gamma(\sum_{k} \alpha_{ijk}) - \sum_{k} \log \Gamma(\alpha_{ijk}) - \log \Gamma(\sum_{k} \tilde{\alpha}_{ijk}) + \sum_{k} \log \Gamma(\tilde{\alpha}_{ijk}) \right]$$

where $\tilde{\alpha}_{ijk} = \alpha_{ijk} + n_{ijk}$. Note that the score decomposes over *i*. One can incorporate structure prior information P(m) as well:

$$score(m) = \log P(\mathcal{D}|m) + \log P(m)$$

Greedy search algorithm: Start with *m*. Consider modifications $m \rightarrow m'$ (edge deletions, additions, reversals). Accept m' if score(m') > score(m). Repeat.

Bayesian inference of model structure: Run MCMC on m.

Bayesian Structural EM for incomplete data

Consider a graphical model with structure *m*, observed data D, hidden variables *X* and parameters θ

The Bayesian score is generally intractable to compute:

score
$$(m) = P(\mathcal{D}|m) = \int \sum_{X} P(X, \theta, \mathcal{D}|m) d\theta$$

Bayesian Structure EM (Friedman, 1998):

- **()** compute MAP parameters $\hat{\theta}$ for current model *m* using EM
- **2** find hidden variable distribution $P(X|\mathcal{D}, \hat{\theta})$
- (3) for a small set of candidate structures compute or approximate

score
$$(m') = \sum_{X} P(X|\mathcal{D}, \hat{\theta}) \log P(\mathcal{D}, X|m')$$

4 $m \leftarrow m'$ with highest score

Directed Graphical Models and Causality

Discovering causal relationships is fundamental to science and cognition.

Although the independence relations are identical, there is a **causal** difference between

- "smoking" \rightarrow "yellow teeth"
- "yellow teeth" \rightarrow "smoking"

Key idea: interventions and the do-calculus:

 $P(S|Y = y) \neq P(S|do(Y = y))$

$$P(Y|S=s) = P(Y|do(S=s))$$

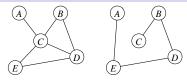
Causal relationships are robust to interventions on the parents.

The key difficulty in learning causal relationships from observational data is the presence of hidden common causes:

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Learning parameters and structure in undirected graphs



 $P(\mathbf{x}|\boldsymbol{\theta}) = \frac{1}{Z(\boldsymbol{\theta})} \prod_{j} g_{j}(\mathbf{x}_{C_{j}}; \boldsymbol{\theta}_{j}) \text{ where } Z(\boldsymbol{\theta}) = \sum_{\mathbf{x}} \prod_{j} g_{j}(\mathbf{x}_{C_{j}}; \boldsymbol{\theta}_{j}).$

Problem: computing $Z(\theta)$ is computationally intractable for general (non-tree-structured) undirected models. Therefore, maximum-likelihood learning of parameters is generally intractable, Bayesian scoring of structures is intractable, etc.

Solutions:

- directly approximate Z(θ) and/or its derivatives (cf. Boltzmann machine learning; contrastive divergence; pseudo-likelihood)
- use approx inference methods (e.g. loopy belief propagation, bounding methods, EP).

(Murray & Ghahramani, 2004; Murray et al, 2006) for Bayesian learning in undirected models.

- Parameter learning in directed models:
 - complete and incomplete data;
 - ML and Bayesian methods
- Structure learning in directed models: complete and incomplete data
- Causality
- Parameter and Structure learning in undirected models