

# Gibbs Sampling for Bayesian Mixture

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# Key concepts

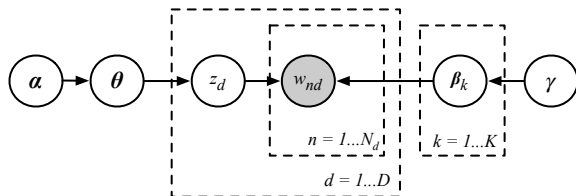
- General Bayesian mixture model
- We derive the Gibbs sampler
- Marginalize out mixing proportions: collapsed Gibbs sampler

# Bayesian document mixture model

Our mixture model has observations  $w_d$  the words in document  $d = 1, \dots, D$ . The parameters are  $\beta_k$  and  $\theta$ , and latent variables  $z$ .

The mixture model has  $K$  components, so the parameters are  $\beta_k, k = 1, \dots, K$ . Each  $\beta_k$  is the parameter of a categorical over possible words, with prior  $p(\beta)$ . The discrete latent variables  $z_d, d = 1, \dots, D$  take on values  $1, \dots, K$ .

Note, that in this model the observations are (the word counts of) entire documents.



$$\begin{aligned}\theta &\sim \text{Dir}(\alpha) \\ \beta_k &\sim \text{Dir}(\gamma) \\ z_d | \theta &\sim \text{Cat}(\cdot) \\ w_{nd} | z_d, \beta &\sim \text{Cat}(f_{z_d})\end{aligned}$$

# Bayesian mixture model

The conditional likelihood is for each observation is

$$p(w_d | z_d = k, \boldsymbol{\beta}) = p(w_d | \boldsymbol{\beta}_k) = p(w_d | \boldsymbol{\beta}_{z_d}),$$

and the prior

$$p(\boldsymbol{\beta}_k) = \text{Dir}(\boldsymbol{\gamma})$$

The categorical latent component assignment probability

$$p(z_d = k | \boldsymbol{\theta}) = \theta_k,$$

with a Dirichlet prior

$$p(\boldsymbol{\theta} | \boldsymbol{\alpha}) = \text{Dir}(\boldsymbol{\alpha}).$$

Therefore, the latent conditional posterior is

$$p(z_d = k | w_d, \boldsymbol{\theta}, \boldsymbol{\beta}) \propto p(z_d = k | \boldsymbol{\theta}) p(w_d | z_d = k, \boldsymbol{\beta}) \propto \theta_k p(w_d | \boldsymbol{\beta}_{z_d}),$$

which is just a discrete distribution with  $K$  possible outcomes.

# Gibbs Sampling

Iteratively, alternately, sample the three types of variables:

Component parameters

$$p(\beta_k | w, z) \propto p(\beta_k) \prod_{d:z_d=k} p(w_d | \beta_k) \propto \text{Dir}(\gamma + c_{mk}),$$

with  $c_{mk} = \sum_{d:z_d=k} c_{md}$ . This is now a categorical model, the mixture aspect having been eliminated.

The posterior latent conditional allocations

$$p(z_d = k | w_d, \theta, \beta) \propto \theta_k p(w_d | \beta_{z_d}),$$

are categorical and mixing proportions

$$p(\theta | z, \alpha) \propto p(\theta | \alpha) p(z | \theta) \propto \text{Dir}(c + \alpha).$$

where  $c_k = \sum_{d:z_d=k} 1$  are the counts for mixture  $k$ .

# Collapsed Gibbs Sampler

The parameters are treated in the same way as before.

If we **marginalize** over  $\theta$

$$\begin{aligned} p(z_d = k | z_{-d}, \alpha) &= \int p(z_d = k | \theta) p(\theta | z_{-d}, \alpha) d\theta \\ &= \int \theta_k p(\theta | z_{-d}, \alpha) d\theta = \frac{\alpha + c_{-d,k}}{\sum_{j=1}^K \alpha + c_{-d,j}}, \end{aligned}$$

where index  $-d$  means *all except*  $d$ , and  $c_k$  are counts; we derived this result when discussing pseudo counts.

The **collapsed** Gibbs sampler for the latent assignments

$$p(z_d = k | w_d, z_{-d}, \beta, \alpha) \propto p(w_d | \beta_k) \frac{\alpha + c_{-d,k}}{\sum_{j=1}^K \alpha + c_{-d,j}},$$

where now all the  $z_d$  variables have become **dependent** (previously they were conditionally independent given  $\theta$ ).

Notice, that the Gibbs sampler exhibits the *rich get richer* property.