## **4F13: Machine Learning**

http://learning.eng.cam.ac.uk/zoubin/ml06/

Department of Engineering, University of Cambridge Michaelmas 2006

#### Lecture 8 Statistical sampling and Monte Carlo

Iain Murray

i.murray+ta@gatsby.ucl.ac.uk

## **Review: probabilistic modelling**

#### Data $\mathcal{D}$ , model $\mathcal{M}$ ; what do we know about x?

Bayesian prediction with unknown parameters  $\theta$ :

$$P(x|\mathcal{D}, \mathcal{M}) = \int P(x, \theta | \mathcal{D}, \mathcal{M}) \, \mathrm{d}\theta \quad \text{Marginalization}$$
$$= \int P(x|\theta, \mathcal{D}, \mathcal{M}) \underbrace{P(\theta | \mathcal{D}, \mathcal{M})}_{\text{from Bayes' rule}} \, \mathrm{d}\theta \quad \text{Product rule}$$

Also marginalize any hidden variables:

$$P(x|\mathcal{D},\mathcal{M}) = \int \sum_{h} P(x,h|\theta,\mathcal{D},\mathcal{M}) \sum_{H} P(\theta,H|\mathcal{D},\mathcal{M}) \, \mathrm{d}\theta$$

### **Regression example**



#### "Review" continued

The **EM** algorithm uses sums or integrals for its sufficient statistics:

$$E[s(h,v)|\theta] = \int s(h,v)P(h|v,\theta) \, \mathrm{d}h$$

Lecture 11 will be on model comparison:

$$P(\mathcal{D}|\mathcal{M}) = \int P(\mathcal{D}|\theta, \mathcal{M}) P(\theta|\mathcal{M}) \, \mathrm{d}\theta$$

Inference only needs mechanical use of marginalization, the product rule and Bayes' rule. . . provided we can do all the sums and integrals

## The trouble with sums

100 binary variables  $x_i \in \{0, 1\}$ , could be:

- assignments of 100 data points in a mixture of 2 Gaussians
- a tiny patch of pixel labels in computer vision
- a *tiny* patch of idealized magnetic iron



There are  $2^{100}$  possible states

The age of the universe  $\approx 2^{98}$  picoseconds

Sum might decompose (e.g. belief propagation) . . . otherwise must approximate

## The trouble with integrals



Only some integrals have analytic solutions Numerical quadrature is feasible in low dimensions (1, 2 or 3)

Multivariate integrals occasionally decompose into standard integrals:

- Linear Gaussian models
- Exponential family models with conjugate priors

Discretization or quadrature are infeasible in high dimensions:

- Only allow each variable to take on 2 settings
- There are  $2^{100}$  possible joint settings. . .

# **Regression example (2)**



## **Statistical sampling**

What is the average height h of people p in Cambridge C?

$$\begin{split} E_{p\in\mathcal{C}}[h(p)] &\equiv \frac{1}{|\mathcal{C}|} \sum_{p\in\mathcal{C}} h(p) \\ &\approx \frac{1}{S} \sum_{s=1}^{S} h(p^{(s)}), \text{ for random survey of } S \text{ people } \{p^{(s)}\} \in \mathcal{C} \end{split}$$

What is the distribution over unknown x?

$$p(x|\mathcal{D}) = \int P(x|\theta, \mathcal{D}) P(\theta|\mathcal{D}) \, \mathrm{d}\theta = E_{P(\theta|\mathcal{D})} [P(x|\theta, \mathcal{D})]$$
$$\approx \frac{1}{S} \sum_{s=1}^{S} P(x|\theta^{(s)}, \mathcal{D}), \quad \theta^{(s)} \sim P(\theta|\mathcal{D})$$

This technique is also known as simple Monte Carlo Estimates are unbiased, variance  $\sim 1/S$  "independent of dimension"

## A dumb approximation of $\pi$

$$P(x,y) = \begin{cases} 1 & 0 < x < 1 \text{ and } 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$
$$\pi = 4 \iint \mathbb{I}\left((x^2 + y^2) < 1\right) P(x,y) \, \mathrm{d}x \, \mathrm{d}y$$

octave:1> N=12; a=rand(N,2); 4\*mean(sum(a.\*a,2)<1)
ans = 3.3333
octave:2> N=1e7; a=rand(N,2); 4\*mean(sum(a.\*a,2)<1)
ans = 3.1418</pre>

## Monte Carlo and Insomnia



**Enrico Fermi** (1901–1954) took great delight in astonishing his colleagues with his remakably accurate predictions of experimental results. . . he revealed that his "guesses" were really derived from the statistical sampling techniques that he used to calculate with whenever insomnia struck in the wee morning hours!

—*The beginning of the Monte Carlo method*, N. Metropolis

## Sampling from distributions

Draw points from the unit area under the curve



Draw probability mass to left of point,  $u \sim \text{Uniform}[0,1]$ 

Sample  $x(u) = c^{-1}(u)$ , where  $c(x) = \int_{-\infty}^{x} P(x') dx'$ 

**Problem:** often can't even normalize P, e.g.  $P(\theta|\mathcal{D}) \propto P(D|\theta)P(\theta)$ Will call unnormalized version  $P^*$  (as in MacKay's textbook).

## **Rejection sampling**

Sampling underneath a  $P^*(x) \propto P(x)$  curve is also valid



 $\begin{array}{l} {\rm Draw \ underneath \ a \ simple} \\ {\rm curve \ } cQ^*(x) \geq P^*(x) {\rm :} \\ {\rm - \ Draw \ } x \sim Q(x) \\ {\rm - \ height \ } h \sim {\rm Uniform}[0,cQ^*(x)] \end{array}$ 

Discard the point if above  $P^*$ , i.e. if  $h > P^*(x)$ 

Computing  $P^*(x)$  and  $Q^*(x)$ , then *throwing* x *away* seems wasteful Instead rewrite the integral as an expectation under Q:

$$\int f(x)P(x) \, \mathrm{d}x = \int f(x)\frac{P(x)}{Q(x)}Q(x) \, \mathrm{d}x, \qquad (Q(x) > 0 \text{ if } P(x) > 0)$$
$$\approx \frac{1}{S}\sum_{s=1}^{S} f(x^{(s)})\frac{P(x^{(s)})}{Q(x^{(s)})}, \quad x^{(s)} \sim Q(x)$$

Unbiased; but light-tailed Q(x) can give the estimator infinite variance . . . and you might not notice.

Importance sampling applies when the integral is not an expectation.

## Importance sampling (2)

Previous slide assumed we could evaluate  $P(x) = P^*(x)/\mathcal{Z}_P$ 

$$\int f(x)P(x) \, \mathrm{d}x \approx \frac{\mathbb{Z}_Q}{\mathbb{Z}_P} \frac{1}{S} \sum_{s=1}^S f(x^{(s)}) \frac{P^*(x^{(s)})}{Q^*(x^{(s)})}, \quad x^{(s)} \sim Q(x)$$
$$\approx \frac{1}{S} \sum_{s=1}^S f(x^{(s)}) \frac{w^{(s)}}{\frac{1}{S} \sum_{s'} w^{(s')}}$$

This estimator is **consistent** but **biased** 

**Exercise:** Prove that  $\mathcal{Z}_P/\mathcal{Z}_Q \approx \frac{1}{S}\sum_s w^{(s)}$ 

## Summary so far

- Sums and integrals, often expectations, occur frequently in statistics
- Monte Carlo approximates expectations with a sample average
- **Rejection sampling** draws samples from fiddly distributions
- Importance sampling applies Monte Carlo to any sum/integral
- If Q(x) is a poor global fit:
  - rejection samplers almost always reject (large c needed)
  - importance sampling is dangerous (large or infinite variance)

In high dimensions finding a good Q(x) is hard. What then?